

Trends in the Covariance Structure
of Earnings in the U.S.:
1969-1987

Robert A. Moffitt
Johns Hopkins University

Peter Gottschalk
Boston College

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Abstract

We examine the increasing variance of earnings of white men over the 1970s and 1980s by focusing on changes in the covariance structure of earnings. Using data from the Michigan PSID from 1969-1987, we find that about half of the increase has arisen from an increase in the variance of the permanent component of earnings and half from an increase in the variance of the transitory component, where the transitory component is composed of serially correlated shocks that die out within three years. We thus find that increases in the variability of earnings are of equal importance to increases in the dispersion of permanent earnings in explaining recent increases in earnings inequality.

Considerable recent attention has been focused on the increase in earnings inequality in the U.S. over the 1970s and 1980s; see Levy and Murnane (1992) for a comprehensive list of the many studies. A growing body of research has shown that inequality in earnings grew over this period not only from an increase in returns to education and experience but also from an increase in inequality within groups of workers of similar age and education. Furthermore, the increase in inequality appears to have occurred throughout the earnings distribution, for the proportion of high-earnings workers as well as of low-earnings workers increased during the 1970s and 1980s. An increase in the dispersion of wage rates has also accompanied that of earnings.

While this literature has firmly established that an increase in the cross-sectional dispersion of earnings and wages has occurred--that is, that the variance of the marginal distribution of earnings and wages has gone up--it has not been established whether the autocovariance structure of earnings and wages--that is, the structure of earnings dynamics--has also shifted. Our study examines this question.

The simplest and most widely known autocovariance structure assumes earnings to be composed of a permanent component and a white-noise transitory component. For this simple model the question we examine is whether the increase in the cross-sectional variance of earnings has arisen from a increased variance of the permanent component or of the transitory component--the latter would imply that there has been an increase in the variability of earnings. However, the literature on earnings dynamics has established that the earnings process is more complex than the simple model implies, containing both random walks and serially-correlated transitory components (Lillard and Willis (1978), Lillard and Weiss (1979), Hause (1977,1980), MaCurdy(1982), Abowd and Card (1989); see Atkinson et al., 1992, for a survey). We therefore examine whether the parameters of such a more complex process have shifted

over time in more general ways.

We also examine trends in earnings mobility. We show that the degree of earnings mobility in an economy is closely related to the autocovariance structure of earnings and, more specially, to the relative magnitudes of its permanent and transitory components. Earnings mobility is positively (negatively) related to the variance of the transitory (permanent) component. Trends in the variances of these two components can therefore be expected to have effects on mobility as well.

We find that the variance of the transitory component of earnings has increased over the 1970s and 1980s in approximately equal magnitude to an increase in the variance of the permanent component. This finding--of a substantial increase in the variability of earnings--has significant implications for research on the causes of increased dispersion, for most hypotheses to date have implicitly presumed the dispersion to have arisen from an increased permanent variance. Changes in the price of human capital (skill) arising from labor demand shifts (e.g., from skill-based technical change as argued by Bound and Johnson, 1992); changes in the dispersion of the quantity of human capital generated by the educational system; increases in the magnitude and dispersion of rents; and other factors all presumably have considerable persistence (see Levy and Murnane for references). Possible explanations for an increased transitory variance are quite different--for example, from an increase in competition both domestic and foreign; the decline of regulation, unions and administered prices; the increase in temporary employment; and change in other factors that increase overall turbulence.

The paper is composed of four sections. The first section briefly discusses the panel data to be used and how it is formulated for estimation of dynamic earnings processes. The second section provides the main results of the paper. The third section extends the results to a discussion of a mobility, while the fourth section takes up several additional issues of importance.

I. Data and Variable Construction

We use the Panel Study on Income Dynamics (PSID), a longitudinal survey which has followed a sample of households from the civilian non-institutional population of the U.S. since 1968. Approximately 5,000 households were interviewed in the initial year of the survey, including a supplementary low-income sample (the SEO) which we also include in our analysis (sample weights are used throughout). Members of the original 1968 households and their offspring have been followed through 1988, the most recent year of data available at the time this analysis was conducted. The primary advantage of the PSID is its long period of coverage and its conformity with cross-sectional measures of inequality.¹ A disadvantage of the PSID is that relatively little information is available on the education and earnings of individuals who are not heads of households.

Following the practice of most previous studies of inequality, we analyze only white males. There are larger sample sizes for whites than for blacks, and the problem of zero earnings is less of a problem for males than for females. We restrict our sample to heads of household 20-59 who had positive hours of work and earnings in the year prior to interview and who were not in school. Our sample--white male heads of household--is likely to have a smaller transitory earnings variance than other groups in the population, and should therefore provide us with a lower bound estimate of earnings variability in the labor force as a whole.

We include every annual observation for each individual for which these restrictions are met; thus the sample is not "continuous" (i.e., there are missing years for some individuals). This permits us to maximize the sample size used for the construction of each element of the covariance matrix.² The earnings and wage measures we examine are the log of real annual earnings (wage and salary only) in the year prior to interview, and the log of real weekly earnings in that year. We exclude the first two years of the survey, 1968 and 1969, because wage and salary earnings data asked in those years were bracketed. Thus our analysis includes the years 1970-1988 and our earnings and

wage measures cover the period 1969–1987. The real figures are obtained by deflating the nominal values by the GNP personal consumption expenditure deflator (base 1982).³ Our final sample includes 2,781 individuals with a total of 25,194 person-year observations.

We conduct most of our work with residuals that are obtained from regressions of these earnings and wage measures on a set of education, age, and year variables; however, we also conduct extra tests using unadjusted earnings and wages. Regressions are estimated separately for each year and by 10-year age interval (20–29, 30–39, 40–49, 50–59), each regression containing education dummies for 0–8, 9–11, 12, 13–15, and 16+ years. We choose this level of disaggregation to maintain a minimum of 250 observations per regression; there are 78 regressions altogether (4 age categories, 19 years).⁴ Using the residuals from these regressions, we calculate earnings and wage variances in each year within each age interval. We also follow individuals forward through time (up to age 59) and compute covariances between their residuals at different ages. We follow different cohorts separately, and thereby obtain variances and covariances at the same age but at different calendar points in time. This will provide the basis for our examination of whether the covariance structure of earnings has shifted.

The covariance matrix pooled over all ages and years has 553 elements, consisting of 76 variances and 477 covariances. The construction of the matrix is elaborate in more detail in Appendix A.

II. Main Results

Simple Permanent-Transitory Models. The simplest and most well-known model of the earnings structure is the canonical permanent-transitory model with white-noise transitory component. Let y_{ia} be the level of earnings (or, in our case, its residual) for individual i at age a . Then

$$y_{ia} = \mu_i + v_{ia} \tag{1}$$

where μ_i is a time-invariant individual component with variance σ_μ^2 and v_{ia} is a serially uncorrelated transitory component with variance σ_v^2 . The variances and covariances in the data can be used in a simple fashion to estimate σ_μ^2 and σ_v^2 because $\text{Var}(y_{ia}) = \sigma_\mu^2 + \sigma_v^2$ and $\text{Cov}(y_{ia}, y_{ia'}) = \sigma_\mu^2$. Hence the permanent variance can be estimated by the autocovariances and the transitory variance can be estimated from the difference between the variances and covariances. The variance of the permanent component is thus synonymous with persistence of earnings. Whether that variance has been increasing over time can be determined by examining whether covariances have been rising; and trends in the transitory variance can be detected by shifts in the difference between variances and covariances.

Another simple method of estimating changes in the permanent variance is to estimate the variance of average earnings (or its residual) over two separate calendar periods. Changes in the transitory variance can be estimated by the change in the variance of the deviations around the two averages. We have conducted an analysis of this type in a prior study (Gottschalk and Moffitt, 1994). However, in addition to requiring the choice of arbitrary calendar time intervals, that method is problematic if the permanent variance does not shift abruptly at interval endpoints but instead trends smoothly upward. In that case, deviations from average earnings during the early and later parts of the interval will be incorrectly attributed to transitory earnings. To instead analyze trends year-by-year requires the approach we are taking.⁵

Figures 1(a)-(d) show trends in variances and in covariances at different lag orders for different age groups.⁶ For all age groups variances have been increasing, consistent with cross-sectional evidence from the CPS. However, the figures also show unmistakable evidence of an increase in covariances as well. The covariance increases are larger for the older age groups and for the low-order covariances, but are positive in almost all cases (though there is some hint of a decline in the final year or two at the older ages). It is also clear from the figures that, although the variance shows

considerable fluctuation--much related to the business cycle--the gap between the variance and the covariances has been growing over time, especially between the variances and the high-order covariances. Because this gap proxies the transitory variance, there is graphical evidence that that variance has increased.⁷

Table 1 provides a regression method of summarizing these patterns, using all 553 individual variances and covariances in our data (see Appendix A). The model in column 1 has only an intercept and a "diagonal" dummy variable (D) equal to 1 if the element is a variance (i.e., it falls on the diagonal of the covariance matrix) and 0 if not, and thus has no time trends (for illustration). The use of the diagonal dummy is convenient for summarizing the difference between variances, which lie on the diagonal, and covariances, which do not. In terms of the canonical permanent-transitory model, therefore, the intercept in column (1) is an estimate of σ_{μ}^2 and the coefficient on D is an estimate of σ_{ν}^2 . The regression shows an average covariance of approximately .13 and an average transitory variance of .18, implying a total variance of approximately .31 and a correlation coefficient of .41, an estimate close to other estimates of random effects earnings models.

The other columns show estimates of how the intercept and slope coefficient have trended over time by including year ("t") in level form and interacted with D, controlling for age effects in the covariances and variances, for distance off the diagonal, and for the unemployment rate. The simplest specification in the second column shows that covariances trended at .0058 per year and that the transitory variance trended at .0059 per year, estimates which are not significantly different from one another. Thus the permanent and transitory variances appear to have trended upward at approximately the same rate. Including the unemployment rate (column 3) shows procyclical variances and covariances but does not affect the variance trends.⁸ The fourth column, which permits the covariances to trend at different rates at different distances off the diagonal, shows strongly what

was apparent in the figures, namely, that the low-order covariances have increased more rapidly than the high-order covariances.

Table 2 allows the coefficients on t and Dt to differ for the periods 1969-1980 and 1981-1987. The most striking finding in the table is the much greater relative growth of the permanent variance in the 1970s and the relatively greater rate of growth of the transitory variance in the 1980s. Indeed, in the 1970s the transitory variance appears to have fallen, or at least not to have changed significantly. This pattern holds both for all the elements of the covariance matrix as well as for the low-order elements shown in the lower portion of the table, where it is also seen that the relatively greater rate of growth of the low-order covariances relative to high-order covariances was concentrated in the 1980s as well. Put differently, the low-order covariances almost doubled from the earlier period to the later one, whereas all covariances (and by implication the high-order ones) grew much less over the same period of time. As we shall show in the next section, this pattern is a result of the same forces that cause the coefficient on Dt (i.e., the "transitory" variance) to rise faster in the 1980s than in the 1970s.

More Realistic Earnings Dynamics Models. Both the Figures and the regressions in Table 1 show that the canonical permanent-transitory model is an incorrect description of the autocovariance structure of earnings. The existing literature on earnings dynamics (see Atkinson et al., 1992 for a survey) finds this to be the case as well. Most importantly, the past literature has found the transitory component to be strongly serially correlated, but arising from an ARMA process of low order (e.g., of order no greater than 2).

Table 3, which displays the covariances and correlations of log annual earnings in our data over all years combined, shows covariance and correlation patterns that are similar to those found in prior work. Covariances and correlations fall rapidly over the first two or three orders and then decline at a much slower rate at higher orders. The covariances and correlations do not fall to zero, but appear to asymptote, consistent with the presence of an

individual effect (i.e., between periods sufficiently far apart, only μ is in common). These patterns would appear to be reasonably well fit by a model with a time-invariant individual effect (to explain the asymptote) plus a low-order ARMA error, the AR component to explain the long declining tail and the other, MA, component to explain the more rapid decline in the first one or two orders.

Another important feature of the autocovariance structure is the notable positive correlation of covariances and correlations with age, holding the order constant. An ARMA model stationary in age is not capable of explaining such a pattern; instead some non-stationary age-related process is required. To capture this pattern we will test for a random walk in age, since random walks generate increasing covariances over time.

Table 4 shows the results of fitting several error components models to the 553 variances and covariances, although, for the moment, not allowing any parameters to vary with calendar time. The models estimated are shown in the notes to the table. The standard errors are computed from the empirical covariance matrix of the residuals (see Appendix B).

The first column fits a model with an individual permanent component and an AR(1) transitory component, while the second column expands the transitory component to an ARMA(1,1). All parameters are significant. The high value of the estimated ρ is consistent with the slow decline in autocorrelations and the large negative estimate for θ implies that the autocorrelations drop off from order 1 to order 2, both as were seen in Table 3. The variance of the individual effect is significant, reflecting the non-zero asymptote. The third and fourth columns of the table attempt to capture the increasing covariances of earnings with age, in one case with a random walk in the individual effect and in the other case by a random growth rate (Hause, 1977, 1980; Lillard and Weiss, 1979). Both models show significant age effects, and the fit is essentially the same for both. The choice is therefore arbitrary in our case, so we pick the random walk specification based on findings from other studies which have used fourth-moments to

distinguish the two.⁹ In other results, we tested ARMA(1,2) and ARMA(2,1) specifications for the transitory effect while maintaining the random-walk specification for the individual effect. In neither case was the fit significantly improved and in neither case was the additional parameter significant. An ARMA(1,1) with a random-walk individual effect hence fits our data adequately.¹⁰

These results are fairly consistent with past work on earnings dynamics. Our model is a bit more refined than early models such as Lillard and Willis (1978), who assumed only an individual effect and an AR(1) transitory effect. But recent, more flexible specifications, such as those tested by MaCurdy (1982) and Abowd and Card (1982), find strong MA components as well as random walk components. Both MaCurdy and Abowd-Card find that an MA(2) specification adequately fits the covariance matrix of earnings differences, for example. Our random-walk-plus-ARMA(1,1) model in levels implies an ARMA(1,2) model in differences, slightly different than MaCurdy and Abowd-Card. But the magnitude of our estimated AR(1) implies a rapid fade-out in differences, so the difference with the differenced MA(2) is not large.¹¹

Our main interest is in allowing the parameters of the process to change with calendar time. To introduce time-varying parameters into the model, we estimate the following specification:

$$y_{iat} = \alpha_t \mu_{iat} + v_{iat} \quad (2)$$

$$\mu_{iat} = \mu_{i,a-1,t-1} + \omega_{iat} \quad (3)$$

$$v_{iat} = \rho_t v_{i,a-1,t-1} + \xi_{iat} + \theta_t \xi_{i,a-1,t-1} \quad (4)$$

Equation (2) shows the log earnings (or earnings residual) of person i at age a in year t to be composed of an individual effect (μ_{iat}) with a time-varying factor loading (α_t), and a transitory effect (v_{iat}). The individual effect could represent latent unobservable human capital whose price (α_t) shifts with calendar time. The individual effect follows a random walk as shown in (3) and the transitory effect follows the ARMA(1,1) process shown in (4). As

conventional in these models, we assume the forcing variables ω_{iat} , ξ_{iat} , and the initial value of the individual effect (μ_{i1t}) to be independently distributed over age and time and w.r.t. each other.

Aside from the variance of the initial individual effect, there are five parameters in the model-- α_t , ρ_t , θ_t , and the variances of ω_{iat} and ξ_{iat} --which, together, determine the pattern of variances and covariances. We permit all five to vary linearly with calendar time (and, subsequently, with year dummies):

$$\alpha_t = 1 + b_1 t \quad (5)$$

$$\rho_t = c_0 + c_1 t \quad (6)$$

$$\theta_t = d_0 + d_1 t \quad (7)$$

$$\text{Var}(\omega_{iat}) = e_0 + e_1 t \quad (8)$$

$$\text{Var}(\xi_{iat}) = f_0 + f_1 t \quad (9)$$

The factor loading α_t is normalized to 1 at $t=0$ (1969 in our data), and we let $\text{Var}(\mu_{i1t}) = \sigma_\mu^2$ to establish the baseline variance of the individual effect.

A "permanent" effect in this model is not permanent in the literal sense since the individual effect is permitted to shift over the life cycle and with calendar time. The distinction between the two components in (2) is, instead, based upon a decomposition of shocks into those that are mean-reverting and those that are not. Our decomposition defines permanent shocks to be those that are non-mean-reverting and transitory shocks to be those that are mean-reverting.

We estimate the model by minimum distance using the form suggested by Chamberlain (1984) for the estimation of covariance structures. The mapping of the model (2)-(9) into the variances and covariances necessary for the estimation is given in Appendix B. Robust standard errors are computed from the empirical covariance matrix of the residuals in the moment equations.

Table 5 shows estimates of the model. Initial testing revealed that the time trend coefficients were significant only for α_t and the variance of ξ_{iat} , so column (1) shows a specification with only these two time effects allowed.

The year coefficient for α_t is .029, implying that its factor loading (or the "price of permanent unobserved human capital") increased by approximately 52 percent over the 18-year period 1969-1987 ($1.52 = 1 + .029 \cdot 18$). Thus the model strongly confirms the existence of an increase in the variance of the permanent component. At the same time, the variance of ξ_{iat} --which is a two-period transitory component--almost doubled over the period, increasing from .117 in 1969 to .207 in 1987 ($.207 = .117 + .005 \cdot 18$). Thus the model also confirms that there was a strong increase in the transitory component. However, because of the presence of the autoregressive process, the increase in the transitory variance persists over time in its effect on the variance of the total transitory component, v_{iat} (see equation (4)). But this effect dies out at the rate ρ^2 , implying, at our estimates of that parameter, that the impact is negligible after three years.

The second column in Table 5 shows that the time trends in the other three parameters of the covariance matrix are insignificant. The magnitude of the trend coefficient for ρ is not trivial, implying an increase from .578 to .722 over the period and hence a strengthening of the low-order covariances and a longer persistence of transitory shocks. However, the large standard error on the coefficient makes this result highly uncertain.¹²

These estimates provide an interpretation of figure 1 and an explanation for the differing rates of growth of low-order and high-order covariances. The estimates imply that covariances of earnings within three years of one another reflect not only the permanent component but also the serially-correlated transitory component. Thus they imply that it is incorrect to associate off-diagonal elements with the permanent variance per se, as the simple canonical model assumes. The more rapid increase in the low-order covariances in Figure 1 (and Table 1) than of the high-order covariances simply reflects the fact that the former captures the increasing transitory variance as well as the increasing permanent variance, whereas the latter reflects only the increasing permanent variance. This also implies that, within the simple permanent-transitory model discussed earlier in this

section, it is the gap between the variance and the high-order covariances--not the low-order covariances--that measures the total transitory variance, and this has clearly risen in Figure 1.

One way of assessing the relative importance of the increase in the variance of the permanent component ($\alpha_t \mu_{iat}$) and the transitory component (v_{iat}) is to calculate what the increase in the total variance would have been from 1969 to 1987 had each parameter increased separately. Table 6 shows the results of such an exercise, obtained by calculating the variance of y_{iat} assuming no change in the parameters from 1969-1987, and by then calculating what the 1987 variance would have been had each of the parameters increased by the magnitudes implied by the coefficients in the second column of Table 5.¹³ The results show that the increase in the permanent variance accounted for approximately 40 percent of the increase in total variance and the increase in the transitory variance accounted for approximately 50 percent, with the remainder accounted for by changes in other parameters.¹⁴ Thus, although the change in the transitory variance accounts for slightly more of the change than that of the permanent variance, the two are roughly equal in importance for practical purposes.¹⁵

The estimates thus far have restricted year effects to a linear trend; yet, at minimum, Table 2 indicated that there may be different trends in the 1970s and 1980s. We therefore reestimate the model in column (1) of Table 5, allowing α_t and $\text{Var}(\xi_{iat})$ to take on different values in each year 1969-1987. As shown in Figure 2, the increase in the two parameters occurred in quite different periods. While the permanent variance grew, on average, through about 1982 or 1983, it leveled off or fell subsequently. The transitory variance, on the other hand, showed essentially no trend until 1980 or 1981, when it began to rise. Although it showed a slight decline after 1984, it was still unambiguously higher in the late 1980s than in the early 1980s, opposite to the pattern for the permanent variance. Thus we find additional evidence indicating relatively higher growth rates of the permanent variance in the 1970s and of the transitory variance in the 1980s.¹⁶

III. Mobility

Mobility, defined as a change in individual ranks within a distribution, is closely related to the covariance structure. For example, an increase in an earnings covariance between any two points in time will necessarily lower mobility because earnings in the two periods are more closely related. However, a stronger statement than this can be made. In Appendix C we show that if earnings follow a joint normal distribution, the probability of a change in individual ranks between any two points in time is a function only of the correlation coefficient between earnings at those two points, and not a function of the absolute levels of either of the variances at the two points in time or the covariance.

The intuition for this result is particularly strong in the canonical permanent-transitory model, where the correlation coefficient between earnings at any two points is equal to the fraction of the variance accounted for by the permanent component, or $\sigma_{\mu}^2 / (\sigma_{\mu}^2 + \sigma_{\nu}^2)$. The degree of mobility in this model thus hinges only on the relative sizes of the permanent and transitory variances. A rise in the permanent variance, which increases the average distance between the earnings of different individuals, lowers the chance of a change in rank; a rise in the transitory variance, on the other hand, makes the chance of a change in rank more likely. But a proportional increase in the permanent and transitory variances has no effect on mobility; the two effects exactly cancel. Therefore, to the extent that the permanent and transitory variances have risen at about the same rate, as suggested by our previous results, this model would show little change in mobility.

We should note that the value of the correlation coefficient in a more realistic model, such as one with serially-correlated transitory components, varies depending upon the distance between the two points under consideration. With serially-correlated but mean-reverting transitory components, correlation coefficients fall with that distance and hence mobility is likely to be greater over longer periods. In addition, if mobility is defined instead on the basis of average earnings over multiple years rather than earnings in a

single year, and if it is a change in the rank of mean earnings that is considered, mobility is likely to be lower since the transitory component is a smaller portion of the total variance when earnings are averaged over multiple years.¹⁷

Since our estimated error components model reported in the last section provides a full accounting for the changes in correlation coefficients (i.e., over different distances and intervals) that have occurred during the 1970s and 1980s, a mobility analysis may at first blush appear redundant; that is, our estimated error components model should by itself determine trends in mobility. However, an examination of transition rates between quantiles of the earnings distribution can provide more detail on whether any changes in mobility have occurred at different parts of the distribution (e.g., at top and bottom). We therefore provide a simple quantile analysis of mobility in this section.

Our mobility analysis uses the same data set and covariance structure as used in the previous analyses except that variance elements are eliminated since they are not relevant to mobility. This leaves us with 477 observations, each of which corresponds to a pair of ages in two particular years. Instead of computing covariances for each such cell, we compute quantile mobility rates using five quantiles (i.e., quintiles).¹⁸

Table 7 shows the year-to-year rates of mobility in the sample between quintiles, pooled over all years and ages. Mobility at the upper and lower quintiles is less than in the middle quintiles.¹⁹ At the upper and lower ends there is an approximate one-third chance of changing rank from one year to the next, as opposed to an approximately fifty-fifty chance for the middle quintiles. The mobility table is also remarkable for its symmetry.

Our interest is, once again, in how these mobility rates have changed over time conditional on age. As we discussed previously, the overall shape of mobility trends should follow those of the covariance analysis closely, but should depend primarily upon trends in the correlation coefficients rather than in the covariances. Figures 3(a)-(b) show the trends in both the

correlation coefficient and the mobility rate between the illustrative ages 35 and 36 ("short") and between 35 and 40 ("long").²⁰ The measure of mobility we use is the sum of the off-diagonal elements in each row of Table 7 (i.e., one minus the probability of staying in the same quintile). This measure is the inverse of what is known as the "immobility ratio" (Atkinson et al., 1992). As expected, the correlation coefficients and mobility rates in both diagrams show an extremely close inverse relationship. The one-year-apart correlation coefficient between ages 35 and 36 shows a slight upward trend in the 1970s but a steeper trend in the 1980s, reflecting the pattern of the transitory variance. Correspondingly, there was very little trend in one-year-apart mobility until the late 1970s, when short-term mobility dropped sharply. The five-year-apart correlation coefficient rose steadily over the late 1970s, albeit with considerable fluctuation, but leveled off in the 1980s; correspondingly, five-year mobility dropped steadily in the 1970s but leveled off in the 1980s. These patterns closely reflect the relative patterns of the transitory and permanent variances discussed previously.

Table 8 shows the results of a regression analysis of the mobility rates for all quintiles, all ages, and lag orders. The first row shows that while there was only a small net decline in overall mobility (over all lag orders), a significant decline in mobility occurred in the top and the bottom two quintiles. The subsequent rows of the table show overall mobility rates consistent with Figure 3, falling significantly only for short-term mobility in the 1980s and only for long-term mobility in the 1970s. However, as in the first row, the trends seem to be concentrated in the upper and lower tails of the distribution. Indeed, for the lowest fifth of earners even short-term mobility declined in the 1970s, which is an indirect indication that the variance of serially-correlated transitory shocks has been increasing for that group over the entire period, not just over the 1980s.

IV. Additional Issues

Weekly Wages and Weeks of Work. An important secondary question is the extent to which the increase in instability in earnings profiles signified by the increase in transitory variances has been a result of increasing instability in wage rates or in employment. The literature on the overall increase in cross-sectional dispersion of earnings indicates that a majority of that increase has arisen from increases in the cross-sectional dispersion of wage rates rather than of weeks of work, hours of work, and employment in general (Levy and Murnane, 1992; Burtless, 1990, Table 7). However, there is no necessary reason for the lesser importance of dispersion in cross-sectional employment measures to follow through for the relative importance of permanent and transitory variances. In fact, the literature on life cycle labor supply analysis and on business cycle fluctuations indicates that employment fluctuates with a greater variance than wages, suggesting that transitory components in employment might be considerably stronger than permanent components.

Figures comparable to figure 1 but with trends in the variances and covariances of log real weekly wages and the log of annual weeks worked (available upon request) indicate that both variances and covariances of log weekly wages rose and that they did so in the same pattern as for log real earnings (i.e., with the same relative patterns for high-order and low-order covariances). Clear evidence of increases in the variances and covariances of weeks of work also appear, although the increases in covariances are much weaker than for earnings or wages. This pattern is consistent with a greater relative importance of transitory factors for weeks worked.

Table 9 shows estimates of several models for log real weekly wages and log annual weeks worked. The descriptive regressions show that the increasing variance of log real weekly wages is equally shared between diagonal and off-diagonal elements, as was the case for annual earnings. However, the coefficients are only approximately two-thirds the magnitude of the Table 2 results, thus confirming a role for increasing dispersion in weeks of work.

This is further confirmed by the results for log and absolute weeks worked in the table, which indicate increasing diagonal and off-diagonal elements but greater relative trends for the diagonal elements. This pattern is also consistent with a greater relative importance of increases in transitory variances for weeks worked. The estimates of the error components models shown in the lower half of the table confirm this and show that increases in transitory variances were particularly marked for weeks of work. However, it should be stressed that transitory variances have increased for real weekly wages as well.

These analyses exclude any consideration of changes in the proportion of the population with no weeks worked at all during the year. Those percentages are relatively small for our sample of prime-age white males but have increased over the period. For those 20-29, for example, the percent without work at all during the year increased from 0.4 percent in 1969 to 2.7 percent in 1987. The corresponding percents for those 30-39, 40-49, and 50-59 are, respectively, 0 to 1.7, 1.8 to 2.2, and 8.0 to 11.3.²¹ Our results thus far already indicate increases in the variance of weeks worked in the worker subsample, and our data indicate even larger increases in that variance when nonworkers are included.

Estimates of the descriptive and error components models for absolute weeks of work inclusive of zeros, comparable to those shown in the last column of Table 9, show stronger trends in the permanent variance and weaker trends in the transitory variance.²² We speculate that an entire year without work may be an indication of a serious wage or employment problem that reflects a permanent condition.

Between-Group Trends. The analysis thus far has been conducted entirely on the residuals from earnings and wage regressions, regressions containing education dummies and estimated separately by year and age interval. An important question is whether our results on the relative importance of trends in the variances of the permanent and transitory components of these within-cell earnings components apply as well to log earnings itself. The answer

depends upon the relative importance of trends in the permanent and transitory variances of the between-cell components, which in our case are the components accounted for by education and age differences in earnings.

There is a much larger literature on trends in education and age differences in earnings than on trends in the within component, the literature showing markedly different trends in both over the 1970s and 1980s for both within and between components (see Levy and Murnane for a review). Our education and age coefficients follow the same general pattern over time as those in the past literature, which have been mainly estimated on the CPS, and therefore we do not present them.²³ Instead, we take a simpler approach to this question by reestimating the models we reported in Section IV on log earnings itself rather than on the regression residuals; the difference in results will be an indirect indication of the importance of trends in the between-group variances. Thus, we work with 553 cells of a covariance matrix of log annual earnings over all years, age groups, and lag orders, constructed as described previously for the regression residuals.

The estimates of the descriptive regressions (not shown) indicate that the permanent variance is considerably more important when the "between" is included. Estimates of average permanent and transitory components for the specification in column (1) of Table 1 are .172 and .185, respectively, implying a correlation coefficient of approximately .48 as opposed to our prior estimate of .41. This is to be expected since education levels in our sample are essentially constant for each individual and, therefore, will mainly contribute to the permanent component of earnings. Estimates of trend coefficients comparable to those in columns (2) and (3) of Table 1 show, moreover, approximately the same coefficients on D_t (.0057 and .0056 in the two columns) but somewhat higher coefficients on D (the average covariance) of .0066 and .0068. This higher value reflects a net increase in educational differentials over the period. Estimates of column (4) for the new covariance matrix reveals, however, the same pattern of greater increases of high-order covariances than low-order covariances as found previously.

In order to contrast the within-group and total results we estimate the error components model on both and simulate the implied permanent and transitory variances. The steady-state values in 1969 and 1987 are shown in Table 10.²⁴ As suggested by the descriptive analysis, the results show a higher level of the permanent variance for total log earnings. In addition, there was a slightly greater rate of increase in the permanent variance when the between is included. However, the magnitudes of the changes induced by including trends in the between are not large, and hence none of our substantive findings (e.g., that upward trends in transitory variance are important) are affected.²⁵

V. Summary and Conclusions

In this paper we have examined the source of the increasing cross-sectional variance of male earnings in the U.S. over the 1970s and 1980s by determining its origins in the covariance structure of earnings. Using data from the Michigan Panel Study of Income Dynamics from 1969-1987 for white males, we find that about half of the increase in variance within education and age groups has arisen from an increase in the variance of the permanent component of earnings and half from an increase in the variance of the transitory component, where the transitory component reflects shocks that die out within three years. We thus find that increases in transitory shocks are of equal importance to increases in the dispersion of permanent earnings in explaining recent increases in earnings inequality. Indeed, the increase in transitory shocks was especially great in the 1980s. Other results show that the increase in transitory shocks appears in weekly wages as well as annual earnings, although even greater in annual weeks of work. We also find that transitory shocks are still very important when trends in the variance across education and age groups are considered.

Our investigation of earnings mobility indicates that mobility changed very little over the period, but with a slight fall in long-term mobility in

the 1970s and a slight fall in short-term mobility in the 1980s, the latter reflecting the increase in short-term covariances arising from a higher variance of serially-correlated transitory shocks. These mobility declines are concentrated in the top and bottom quintiles of the earnings distribution.

Our study has been largely a statistical accounting exercise aimed at determining the relative contributions of different error component variances to the upward trend in overall cross-sectional variances, rather than a search for causes. We have conducted a rudimentary exploration of the latter type in Gottschalk and Moffitt (1994), where we found that while some of the increasing transitory variance is a result of a decline in unionization (union jobs have lower transitory variances) and industry shifts, these do not provide a sufficient explanation by themselves--transitory variances have increased within unionized and non-unionized jobs, within all industrial sectors, and even for workers who have stayed with the same firm for up to 10 years. Further work in exploring the sources of increased variability would therefore appear warranted.

Appendix A

Structure of the Autocovariance Matrix

The cells used to construct the autocovariance matrix from the data are broken out separately by age and year, in order to examine trends in the matrix elements over time but holding age constant. The only difficulty in constructing the matrix is the necessity to group the data into age intervals. As noted in the text, we group the data from age 20 through 59 into 4 ten-year age groups (20-29, 30-39, 40-49, 50-59). In each year t of the data, we follow the individuals in each of these four groups through to year $t+1$, year $t+2$, etc. until either the end of our data is reached (1988) or until the age interval in question reaches beyond age 59 (e.g., the 40-49 cohort in 1969 can be followed through to 1979, when the individuals are 50-59, but no further²⁶). Covariances are then calculated between the initial year, t , and each subsequent year.²⁷ A fresh set of cohorts is begun in each year, starting in 1969, and continuing through 1987; the four cohorts (i.e., age groups) started in each year are again followed over time.

Table A-1 shows the cells of the data. The ages shown are in the midpoints of the intervals ("25" for 20-29, "26" for 21-30, etc.). The beginning of the table showed the variances of earnings in the 20-29 group for each of the years 1969 through 1987. The next set of rows shows the autocovariances of earnings between the individuals 20-29 in the initial year and 21-30 ("26") in the following year, computed separately for each year pair 1969-1970 through 1986-1987. The next set of rows shows the autocovariances of earnings of individuals 20-29 in the initial year and 22-31 two years later, which are observed for year pairs 1969-1971 through 1985-1987. Thereafter the covariances between all future years for the cohorts begun in each year. There are 553 cells in total.

This method of grouping ensures that every individual variance and

covariance in the panel is included uniquely in one cell. There are many alternative methods of grouping the data and computing the autocovariance matrix, but there is no reason for any one to be preferred to another except for convenience. Our method is designed to make trends over time, holding age constant, particularly easy to discern.

Table A-1

Elements of the Autocovariance Matrix for Log Annual Earnings

Lower Age	Upper Age	Lower Year	Upper Year	Covariance	Correlation
25	25	69	69	0.25238188	1.00000000
25	25	70	70	0.27661610	1.00000000
25	25	71	71	0.29885597	1.00000000
25	25	72	72	0.28129703	1.00000000
25	25	73	73	0.24614233	1.00000000
25	25	74	74	0.26469568	1.00000000
25	25	75	75	0.43190508	1.00000000
25	25	76	76	0.26104279	1.00000000
25	25	77	77	0.26704491	1.00000000
25	25	78	78	0.30346186	1.00000000
25	25	79	79	0.38524288	1.00000000
25	25	80	80	0.33242864	1.00000000
25	25	81	81	0.39948014	1.00000000
25	25	82	82	0.50573263	1.00000000
25	25	83	83	0.39997789	1.00000000
25	25	84	84	0.31458814	1.00000000
25	25	85	85	0.42641570	1.00000000
25	25	86	86	0.43640994	1.00000000
25	25	87	87	0.42371208	1.00000000
25	26	69	70	0.11633512	0.63167875
25	26	70	71	0.13821284	0.59565863
25	26	71	72	0.11342464	0.52259787
25	26	72	73	0.14004715	0.63837205
25	26	73	74	0.11546026	0.59151964
25	26	74	75	0.12219645	0.58660671
25	26	75	76	0.15682258	0.60795617
25	26	76	77	0.13151195	0.66666760
25	26	77	78	0.14674520	0.65900616
25	26	78	79	0.16755935	0.65699472
25	26	79	80	0.20299283	0.72175769
25	26	80	81	0.19808092	0.68160385
25	26	81	82	0.18401466	0.50621530
25	26	82	83	0.20191927	0.54819073
25	26	83	84	0.16691218	0.63894878
25	26	84	85	0.19559662	0.67716967
25	26	85	86	0.21750304	0.64041851
25	26	86	87	0.21225842	0.61792656
25	27	69	71	0.09866807	0.48208527
25	27	70	72	0.09279338	0.40253686
25	27	71	73	0.09156946	0.47025502
25	27	72	74	0.10232635	0.49919795
25	27	73	75	0.09184921	0.44692778
25	27	74	76	0.10757963	0.47142084
25	27	75	77	0.10857410	0.49677680
25	27	76	78	0.12459753	0.54874119
25	27	77	79	0.13440083	0.57530542
25	27	78	80	0.11694031	0.48949713
25	27	79	81	0.17531619	0.61783509
25	27	80	82	0.16787270	0.49590720

Table A-1 (cont'd)

25	27	81	83	0.16025237	0.51150350
25	27	82	84	0.16942197	0.50655937
25	27	83	85	0.16584386	0.52905995
25	27	84	86	0.15056479	0.53802160
25	27	85	87	0.17926115	0.56929064
25	28	69	72	0.07961085	0.42257982
25	28	70	73	0.08624263	0.38756770
25	28	71	74	0.10617399	0.47172091
25	28	72	75	0.09528073	0.40358414
25	28	73	76	0.09396649	0.43700238
25	28	74	77	0.07332420	0.34859434
25	28	75	78	0.11595002	0.41771346
25	28	76	79	0.08888380	0.45289509
25	28	77	80	0.11372574	0.50108722
25	28	78	81	0.10697949	0.39753841
25	28	79	82	0.15448279	0.39266133
25	28	80	83	0.13684763	0.48681720
25	28	81	84	0.13347704	0.48255631
25	28	82	85	0.15756926	0.42125528
25	28	83	86	0.14326366	0.43837657
25	28	84	87	0.14502792	0.52028730
25	29	69	73	0.07543691	0.37479200
25	29	70	74	0.05999891	0.29206800
25	29	71	75	0.11876027	0.46597082
25	29	72	76	0.10376712	0.42085803
25	29	73	77	0.07791934	0.39243477
25	29	74	78	0.06333104	0.30477997
25	29	75	79	0.10030708	0.36463545
25	29	76	80	0.09770953	0.45692784
25	29	77	81	0.08504214	0.37234993
25	29	78	82	0.14263073	0.40694904
25	29	79	83	0.11128236	0.34348803
25	29	80	84	0.12104363	0.48395931
25	29	81	85	0.15265122	0.42626753
25	29	82	86	0.14965524	0.33430411
25	29	83	87	0.11615068	0.41339401
25	30	69	74	0.05852007	0.28428150
25	30	70	75	0.07869137	0.38110584
25	30	71	76	0.11547119	0.42340466
25	30	72	77	0.07950265	0.35977812
25	30	73	78	0.04002073	0.21406528
25	30	74	79	0.05501892	0.23861011
25	30	75	80	0.10003192	0.35299469
25	30	76	81	0.08818811	0.36814051
25	30	77	82	0.13497743	0.50148070
25	30	78	83	0.14750095	0.42098171
25	30	79	84	0.12826679	0.43084562
25	30	80	85	0.10694778	0.37990074
25	30	81	86	0.15954503	0.46707946
25	30	82	87	0.19699868	0.42347796
25	31	69	75	0.07785410	0.36591022
25	31	70	76	0.06443509	0.32312146
25	31	71	77	0.06429405	0.29175863
25	31	72	78	0.07307886	0.38113335
25	31	73	79	0.05034927	0.25201528
25	31	74	80	0.06154955	0.29281806
25	31	75	81	0.07181829	0.22511274
25	31	76	82	0.08387374	0.37335981

Table A-1 (cont'd)

25	31	77	83	0.10869648	0.34439521
25	31	78	84	0.12169507	0.42629244
25	31	79	85	0.09761055	0.31206388
25	31	80	86	0.13714024	0.47061075
25	31	81	87	0.10243348	0.36542198
25	32	69	76	0.09471803	0.41971012
25	32	70	77	0.05070792	0.22384846
25	32	71	78	0.04703889	0.23764678
25	32	72	79	0.06941981	0.32537881
25	32	73	80	0.05873525	0.29395281
25	32	74	81	0.05499573	0.23045075
25	32	75	82	0.10716559	0.43277929
25	32	76	83	0.08582790	0.27044687
25	32	77	84	0.09595455	0.36425277
25	32	78	85	0.12242578	0.40690334
25	32	79	86	0.10859377	0.33586166
25	32	80	87	0.11165113	0.40749023
25	33	69	77	0.08103649	0.36672029
25	33	70	78	0.06063010	0.31175362
25	33	71	79	0.06243404	0.29113196
25	33	72	80	0.10182218	0.46635152
25	33	73	81	0.06337081	0.26750401
25	33	74	82	0.08950424	0.42171494
25	33	75	83	0.10376041	0.28737979
25	33	76	84	0.08254653	0.31776590
25	33	77	85	0.11165876	0.38523698
25	33	78	86	0.15237092	0.46245530
25	33	79	87	0.11451762	0.40080550
25	34	69	78	0.08296506	0.41443325
25	34	70	79	0.06167349	0.32375240
25	34	71	80	0.07751481	0.34202613
25	34	72	81	0.09557326	0.39507951
25	34	73	82	0.07879655	0.31880094
25	34	74	83	0.07418957	0.23843042
25	34	75	84	0.09042044	0.26223610
25	34	76	85	0.07505823	0.25648842
25	34	77	86	0.09404648	0.33094745
25	34	78	87	0.12590517	0.45942639
25	35	69	79	0.06543087	0.33224462
25	35	70	80	0.08361872	0.35077237
25	35	71	81	0.10130704	0.40228552
25	35	72	82	0.11177885	0.49345778
25	35	73	83	0.09474978	0.33424135
25	35	74	84	0.06696919	0.25255336
25	35	75	85	0.09963772	0.28460584
25	35	76	86	0.07984423	0.29585403
25	35	77	87	0.09608901	0.38417229
25	36	69	80	0.08543528	0.37546050
25	36	70	81	0.07692504	0.28490531
25	36	71	82	0.09017109	0.43294206
25	36	72	83	0.11494263	0.31571937
25	36	73	84	0.05657153	0.23756236
25	36	74	85	0.04857361	0.17567050
25	36	75	86	0.07106139	0.19493965
25	36	76	87	0.06911677	0.27977718
25	37	69	81	0.06430834	0.28440934
25	37	70	82	0.07953987	0.35822704
25	37	71	83	0.11632399	0.33113468

Table A-1 (cont'd)

25	37	72	84	0.09623319	0.30468022
25	37	73	85	0.10688008	0.31351192
25	37	74	86	0.04394220	0.14786917
25	37	75	87	0.09772782	0.32263907
25	38	69	82	0.06382157	0.28097040
25	38	70	83	0.07748352	0.27928911
25	38	71	84	0.10867363	0.31895856
25	38	72	85	0.14270788	0.37922260
25	38	73	86	0.05545543	0.18790020
25	38	74	87	0.06369242	0.26578764
25	39	69	83	0.07637790	0.30371474
25	39	70	84	0.06046641	0.21564954
25	39	71	85	0.11885353	0.33141715
25	39	72	86	0.07812950	0.23960532
25	39	73	87	0.05905213	0.24421106
25	40	69	84	0.10017744	0.36831807
25	40	70	85	0.09822473	0.26115192
25	40	71	86	0.10582135	0.33492992
25	40	72	87	0.07944614	0.30528294
25	41	69	85	0.07648787	0.22863580
25	41	70	86	0.05921662	0.23618920
25	41	71	87	0.11893127	0.42156702
25	42	69	86	0.08850931	0.28819706
25	42	70	87	0.08503483	0.34087897
25	43	69	87	0.10043580	0.36543209
35	35	69	69	0.29575627	1.00000000
35	35	70	70	0.29373972	1.00000000
35	35	71	71	0.23090994	1.00000000
35	35	72	72	0.19957159	1.00000000
35	35	73	73	0.18023090	1.00000000
35	35	74	74	0.23808636	1.00000000
35	35	75	75	0.23458647	1.00000000
35	35	76	76	0.34440097	1.00000000
35	35	77	77	0.23282476	1.00000000
35	35	78	78	0.21873077	1.00000000
35	35	79	79	0.20391977	1.00000000
35	35	80	80	0.22575833	1.00000000
35	35	81	81	0.26517069	1.00000000
35	35	82	82	0.32012594	1.00000000
35	35	83	83	0.38348632	1.00000000
35	35	84	84	0.33689053	1.00000000
35	35	85	85	0.39398481	1.00000000
35	35	86	86	0.42165217	1.00000000
35	35	87	87	0.33897801	1.00000000
35	36	69	70	0.10799405	0.62738270
35	36	70	71	0.11873585	0.64975896
35	36	71	72	0.13604364	0.77529695
35	36	72	73	0.12144991	0.74850054
35	36	73	74	0.12894967	0.75818783
35	36	74	75	0.14176807	0.69900552
35	36	75	76	0.13050343	0.67368649
35	36	76	77	0.16464134	0.71879334
35	36	77	78	0.11531758	0.71537523
35	36	78	79	0.12927109	0.75750547
35	36	79	80	0.14554272	0.83374069
35	36	80	81	0.13293863	0.70135582
35	36	81	82	0.19851673	0.74479164
35	36	82	83	0.19998343	0.67991527

Table A-1 (cont'd)

35	36	83	84	0.17502639	0.72943404
35	36	84	85	0.22216544	0.76894690
35	36	85	86	0.25730562	0.77270412
35	36	86	87	0.22992805	0.79621907
35	37	69	71	0.11326541	0.58458088
35	37	70	72	0.10775334	0.60624414
35	37	71	73	0.11308605	0.72981096
35	37	72	74	0.09850006	0.60809306
35	37	73	75	0.11624987	0.66477877
35	37	74	76	0.13438354	0.59479556
35	37	75	77	0.12446262	0.60690191
35	37	76	78	0.12175314	0.63027885
35	37	77	79	0.10919921	0.63408922
35	37	78	80	0.11489002	0.70376926
35	37	79	81	0.13080792	0.67157580
35	37	80	82	0.14594004	0.60538630
35	37	81	83	0.14802777	0.59296476
35	37	82	84	0.18385733	0.54007643
35	37	83	85	0.19080162	0.66990303
35	37	84	86	0.20717688	0.69794514
35	37	85	87	0.19602736	0.70958374
35	38	69	72	0.10002642	0.51154300
35	38	70	73	0.08640619	0.56832930
35	38	71	74	0.10096207	0.56665153
35	38	72	75	0.08329954	0.54106813
35	38	73	76	0.09783600	0.52078480
35	38	74	77	0.10811593	0.61893031
35	38	75	78	0.09036375	0.55967314
35	38	76	79	0.11673417	0.55901296
35	38	77	80	0.13581815	0.69706610
35	38	78	81	0.12584093	0.60538425
35	38	79	82	0.11600084	0.52903474
35	38	80	83	0.12331923	0.59596576
35	38	81	84	0.13009112	0.56713480
35	38	82	85	0.18628058	0.54793500
35	38	83	86	0.19491268	0.65958365
35	38	84	87	0.18369060	0.68003395
35	39	69	73	0.07755548	0.50012662
35	39	70	74	0.09690851	0.47449027
35	39	71	75	0.10168147	0.55483884
35	39	72	76	0.08385627	0.42354181
35	39	73	77	0.08919678	0.56513403
35	39	74	78	0.11603696	0.58675609
35	39	75	79	0.10149781	0.53643158
35	39	76	80	0.17218863	0.69409284
35	39	77	81	0.12327679	0.58143160
35	39	78	82	0.10227251	0.52635389
35	39	79	83	0.14048602	0.60455185
35	39	80	84	0.11636288	0.53672026
35	39	81	85	0.15313519	0.53239444
35	39	82	86	0.15704782	0.53300832
35	39	83	87	0.17248239	0.62100705
35	40	69	74	0.08112916	0.37933330
35	40	70	75	0.11984755	0.53736040
35	40	71	76	0.11150620	0.43347122
35	40	72	77	0.08898410	0.51808885
35	40	73	78	0.08754521	0.54220403
35	40	74	79	0.09002747	0.49557796

Table A-1 (cont'd)

35	40	75	80	0.12404980	0.60337486
35	40	76	81	0.14039209	0.60424036
35	40	77	82	0.11882276	0.61461105
35	40	78	83	0.12693847	0.54609357
35	40	79	84	0.09475936	0.44967260
35	40	80	85	0.12176836	0.47993331
35	40	81	86	0.13116267	0.57826009
35	40	82	87	0.16255441	0.52122396
35	41	69	75	0.05651727	0.25158133
35	41	70	76	0.10825400	0.38012324
35	41	71	77	0.10597530	0.54365253
35	41	72	78	0.10430018	0.53364146
35	41	73	79	0.10608943	0.56844010
35	41	74	80	0.11702395	0.61962651
35	41	75	81	0.14037138	0.59988745
35	41	76	82	0.19139425	0.58485586
35	41	77	83	0.12554693	0.55712686
35	41	78	84	0.10569432	0.45149954
35	41	79	85	0.13433820	0.50996248
35	41	80	86	0.09627442	0.43771926
35	41	81	87	0.12103767	0.52310531
35	42	69	76	0.11610103	0.51252548
35	42	70	77	0.10175541	0.49230691
35	42	71	78	0.11040124	0.57061544
35	42	72	79	0.09094121	0.48584131
35	42	73	80	0.10432545	0.61240867
35	42	74	81	0.11982831	0.55731284
35	42	75	82	0.16925654	0.57168687
35	42	76	83	0.18662131	0.63410442
35	42	77	84	0.13155154	0.50626558
35	42	78	85	0.11069693	0.44023007
35	42	79	86	0.11391833	0.49777507
35	42	80	87	0.10697071	0.51559679
35	43	69	77	0.11597465	0.48248416
35	43	70	78	0.09771238	0.43497969
35	43	71	79	0.11573113	0.52846091
35	43	72	80	0.08904726	0.46565233
35	43	73	81	0.10974553	0.49809136
35	43	74	82	0.12218862	0.50318561
35	43	75	83	0.15370061	0.57952494
35	43	76	84	0.15568707	0.54548330
35	43	77	85	0.12849729	0.45370867
35	43	78	86	0.09855969	0.47362676
35	43	79	87	0.11581614	0.54499438
35	44	69	78	0.07731330	0.34285575
35	44	70	79	0.09092584	0.43150032
35	44	71	80	0.10954453	0.51532161
35	44	72	81	0.10342399	0.43460711
35	44	73	82	0.16680586	0.57790590
35	44	74	83	0.12880958	0.49143840
35	44	75	84	0.13706969	0.54201022
35	44	76	85	0.14758292	0.46770053
35	44	77	86	0.10594947	0.41480395
35	44	78	87	0.09936566	0.49410779
35	45	69	79	0.07930485	0.35734477
35	45	70	80	0.10178299	0.48580665
35	45	71	81	0.12228159	0.51510462
35	45	72	82	0.09430867	0.33026370

Table A-1 (cont'd)

35	45	73	83	0.11164080	0.51196058
35	45	74	84	0.14807400	0.59480858
35	45	75	85	0.14520793	0.56503630
35	45	76	86	0.12132591	0.42830640
35	45	77	87	0.08282760	0.40402788
35	46	69	80	0.08614411	0.38205010
35	46	70	81	0.11827860	0.53585980
35	46	71	82	0.20187612	0.59900647
35	46	72	83	0.08080180	0.35022378
35	46	73	84	0.13069770	0.60935116
35	46	74	85	0.13187763	0.49723125
35	46	75	86	0.10593031	0.44907824
35	46	76	87	0.10969483	0.43622209
35	47	69	81	0.09974853	0.43268432
35	47	70	82	0.15902232	0.56729410
35	47	71	83	0.11843594	0.48975207
35	47	72	84	0.07165584	0.39273026
35	47	73	85	0.12392844	0.52648532
35	47	74	86	0.10948550	0.50505418
35	47	75	87	0.09193293	0.42156245
35	48	69	82	0.12698106	0.41563058
35	48	70	83	0.11753844	0.47656642
35	48	71	84	0.09405469	0.47117984
35	48	72	85	0.08240910	0.41220891
35	48	73	86	0.10413944	0.44080374
35	48	74	87	0.09303585	0.49249451
35	49	69	83	0.10091445	0.39934180
35	49	70	84	0.09015872	0.45513756
35	49	71	85	0.11079876	0.47588498
35	49	72	86	0.07164230	0.34317330
35	49	73	87	0.10072712	0.50871371
35	50	69	84	0.10108034	0.47189658
35	50	70	85	0.10439749	0.47644525
35	50	71	86	0.09567998	0.45294817
35	50	72	87	0.08599539	0.40434647
35	51	69	85	0.10383210	0.43212285
35	51	70	86	0.08846481	0.45288614
35	51	71	87	0.11258039	0.51408156
35	52	69	86	0.09716044	0.44397917
35	52	70	87	0.11740437	0.48110302
35	53	69	87	0.11416182	0.41074034
45	45	69	69	0.18450732	1.00000000
45	45	70	70	0.18762559	1.00000000
45	45	71	71	0.21260064	1.00000000
45	45	72	72	0.19622150	1.00000000
45	45	73	73	0.15778624	1.00000000
45	45	74	74	0.18656541	1.00000000
45	45	75	75	0.26861350	1.00000000
45	45	76	76	0.24174888	1.00000000
45	45	77	77	0.23420209	1.00000000
45	45	78	78	0.26441058	1.00000000
45	45	79	79	0.28745096	1.00000000
45	45	80	80	0.28096509	1.00000000
45	45	81	81	0.33564005	1.00000000
45	45	82	82	0.45541647	1.00000000
45	45	83	83	0.38459008	1.00000000
45	45	84	84	0.37185514	1.00000000

Table A-1 (cont'd)

45	45	85	85	0.42152591	1.00000000
45	45	86	86	0.34952142	1.00000000
45	45	87	87	0.28788574	1.00000000
45	46	69	70	0.12415186	0.79348734
45	46	70	71	0.13166740	0.81163386
45	46	71	72	0.12863474	0.82304806
45	46	72	73	0.13862347	0.87038045
45	46	73	74	0.13570983	0.86070611
45	46	74	75	0.14612226	0.78290837
45	46	75	76	0.17768590	0.83083201
45	46	76	77	0.14029805	0.75754916
45	46	77	78	0.16328858	0.73217632
45	46	78	79	0.18280427	0.76025639
45	46	79	80	0.22136823	0.81324685
45	46	80	81	0.22921152	0.88526701
45	46	81	82	0.28535136	0.81286425
45	46	82	83	0.31366553	0.77304395
45	46	83	84	0.20215241	0.77754417
45	46	84	85	0.24213282	0.82924466
45	46	85	86	0.23775723	0.77618131
45	46	86	87	0.21223617	0.86042142
45	47	69	71	0.12848324	0.78067712
45	47	70	72	0.12317067	0.77336302
45	47	71	73	0.12089597	0.72752448
45	47	72	74	0.12862652	0.78044227
45	47	73	75	0.12587253	0.71461933
45	47	74	76	0.14562659	0.72551218
45	47	75	77	0.14434890	0.78803348
45	47	76	78	0.15121667	0.75420824
45	47	77	79	0.15835363	0.74282666
45	47	78	80	0.16979023	0.74665160
45	47	79	81	0.18792305	0.75879846
45	47	80	82	0.21245219	0.64661513
45	47	81	83	0.23838206	0.69135480
45	47	82	84	0.18394933	0.69626621
45	47	83	85	0.21366987	0.69386177
45	47	84	86	0.23437112	0.77185127
45	47	85	87	0.21680244	0.71730719
45	48	69	72	0.12492398	0.74455867
45	48	70	73	0.12866905	0.75795486
45	48	71	74	0.12590696	0.71727696
45	48	72	75	0.13379141	0.70806834
45	48	73	76	0.12666606	0.69194649
45	48	74	77	0.13553801	0.73665650
45	48	75	78	0.13910216	0.77403928
45	48	76	79	0.13870509	0.73082841
45	48	77	80	0.14330316	0.67370814
45	48	78	81	0.15775888	0.68453991
45	48	79	82	0.19335256	0.59734172
45	48	80	83	0.23128612	0.70558171
45	48	81	84	0.17316764	0.69753670
45	48	82	85	0.25492586	0.66752297
45	48	83	86	0.18188307	0.54399355
45	48	84	87	0.18189091	0.73814206
45	49	69	73	0.12112981	0.73876290
45	49	70	74	0.11065046	0.68589170
45	49	71	75	0.12782530	0.59630672
45	49	72	76	0.12983524	0.69034977

Table A-1 (cont'd)

45	49	73	77	0.12236924	0.72040496
45	49	74	78	0.13129965	0.66002213
45	49	75	79	0.14352933	0.73242269
45	49	76	80	0.13613068	0.65258827
45	49	77	81	0.17348321	0.73389920
45	49	78	82	0.18967184	0.57297328
45	49	79	83	0.24327555	0.75832928
45	49	80	84	0.18120065	0.72333408
45	49	81	85	0.18940768	0.65303267
45	49	82	86	0.15052001	0.48627933
45	49	83	87	0.15163712	0.52937859
45	50	69	74	0.10595853	0.69689056
45	50	70	75	0.10740767	0.59456272
45	50	71	76	0.12509121	0.62695063
45	50	72	77	0.12764210	0.69993807
45	50	73	78	0.11146100	0.62089458
45	50	74	79	0.13645481	0.72100349
45	50	75	80	0.14015426	0.71166312
45	50	76	81	0.14228485	0.68174878
45	50	77	82	0.27442679	0.72990877
45	50	78	83	0.16644934	0.61055788
45	50	79	84	0.19195626	0.71259508
45	50	80	85	0.21580317	0.73455694
45	50	81	86	0.18837655	0.67068228
45	50	82	87	0.18004192	0.56108297
45	51	69	75	0.11017644	0.58582978
45	51	70	76	0.10598551	0.62950711
45	51	71	77	0.10387774	0.56799145
45	51	72	78	0.12982331	0.67451179
45	51	73	79	0.12190785	0.71636895
45	51	74	80	0.11943817	0.64809972
45	51	75	81	0.13743028	0.67526210
45	51	76	82	0.24390376	0.65324241
45	51	77	83	0.21548279	0.71901995
45	51	78	84	0.13898516	0.63751855
45	51	79	85	0.21833860	0.72082696
45	51	80	86	0.21576172	0.69664288
45	51	81	87	0.16075924	0.55834693
45	52	69	76	0.11258549	0.51728146
45	52	70	77	0.09897096	0.57823989
45	52	71	78	0.11500808	0.60762162
45	52	72	79	0.12980932	0.68059063
45	52	73	80	0.10820331	0.62321354
45	52	74	81	0.11690633	0.63276818
45	52	75	82	0.17534886	0.60304985
45	52	76	83	0.17508725	0.67606120
45	52	77	84	0.18123254	0.63847351
45	52	78	85	0.15021171	0.60967338
45	52	79	86	0.18861341	0.66028879
45	52	80	87	0.20414454	0.62637655
45	53	69	77	0.10026401	0.52977759
45	53	70	78	0.09789989	0.56913478
45	53	71	79	0.11192195	0.61013270
45	53	72	80	0.10823988	0.54873668
45	53	73	81	0.10266486	0.58421503
45	53	74	82	0.14038564	0.42919335
45	53	75	83	0.15006270	0.61103059
45	53	76	84	0.15129720	0.61150573

Table A-1 (cont'd)

45	53	77	85	0.17369149	0.59869221
45	53	78	86	0.13227547	0.60695741
45	53	79	87	0.18631318	0.59992570
45	54	69	78	0.09505115	0.52934985
45	54	70	79	0.11191512	0.64250813
45	54	71	80	0.09770058	0.54895333
45	54	72	81	0.10575867	0.51437867
45	54	73	82	0.18349166	0.60220801
45	54	74	83	0.11290932	0.50741697
45	54	75	84	0.13213210	0.56788281
45	54	76	85	0.14440478	0.44198949
45	54	77	86	0.16777826	0.57943321
45	54	78	87	0.12792258	0.53694397
45	55	69	79	0.11060217	0.61021201
45	55	70	80	0.10992745	0.61344591
45	55	71	81	0.09653611	0.49851443
45	55	72	82	0.13830385	0.36875467
45	55	73	83	0.10459329	0.45267594
45	55	74	84	0.12097583	0.47870097
45	55	75	85	0.11828959	0.38711642
45	55	76	86	0.13427016	0.47481047
45	55	77	87	0.14171002	0.51670066
55	55	69	69	0.17728938	1.00000000
55	55	70	70	0.30183911	1.00000000
55	55	71	71	0.29061541	1.00000000
55	55	72	72	0.28286051	1.00000000
55	55	73	73	0.24741125	1.00000000
55	55	74	74	0.18012075	1.00000000
55	55	75	75	0.19401375	1.00000000
55	55	76	76	0.27513949	1.00000000
55	55	77	77	0.25312508	1.00000000
55	55	78	78	0.23749102	1.00000000
55	55	79	79	0.26942362	1.00000000
55	55	80	80	0.24637073	1.00000000
55	55	81	81	0.28905820	1.00000000
55	55	82	82	0.54746327	1.00000000
55	55	83	83	0.45761974	1.00000000
55	55	84	84	0.37930624	1.00000000
55	55	85	85	0.55924920	1.00000000
55	55	86	86	0.46763581	1.00000000
55	55	87	87	0.45397922	1.00000000

Appendix B

The Minimum Distance Method and the Mapping of the Covariance Model

Minimum Distance Method. Let $s_{im} = y_{ij}y_{ik}$, where y_{ij} and y_{ik} are the log earnings (or residuals) for individual i for age-year "locations" j and k , and where $m=1, \dots, M$ is the moment generated by the product of residuals at locations j and j' . In our case, $M=553$. Posit the model

$$s_{im} = f(\theta, j, k) + \epsilon_{im} \quad i=1, \dots, N, \quad m=1, \dots, M \quad (B1)$$

where θ is a $L \times 1$ vector of parameters. Then the set of M equations in (B1) constitutes an SUR system whose efficient estimation requires an initial consistent estimate of the covariance matrix of the ϵ_{im} . However, following the findings and recommendations of Altonji and Segal (1991) on bias in estimating covariance structures of this type, we employ the identity matrix for the estimation.²⁸ Hence we choose θ to minimize the sum of squared residuals:

$$\text{Min}_{\theta} \sum_{i=1}^N \sum_{m=1}^M [s_{im} - f(\theta, j, k)]^2 \quad (B2)$$

or, equivalently, since f is not a function of i ,

$$\text{Min}_{\theta} \sum_{m=1}^M [\bar{s}_m - f(\theta, j, k)]^2 \quad (B3)$$

where \bar{s}_m is the mean (over i) of s_{im} (i.e., a covariance).

To obtain standard errors, we apply the extension of Eicker-White methods in the manner suggested by Chamberlain, using the residuals from (B1), each of which we denote e_{im} . Let Ω be the $M \times M$ covariance matrix of the e_{im} , each element of which is estimated by:²⁹

$$\hat{\sigma}_{mm'} = (1/N) \sum_{i=1}^N e_{im} e_{im'} \quad (B4)$$

Let the NMxNM covariance matrix of individual residuals be

$$\Delta = \begin{bmatrix} \Omega & & & & \\ & \Omega & & 0 & \\ & & \Omega & & \\ & 0 & & \Omega & \\ & & & & \Omega \end{bmatrix}$$

Then

$$\text{Cov}(\hat{\theta}) = (G'G)^{-1} G' \Delta G (G'G)^{-1} \quad (B5)$$

where G is the NMxL matrix of gradients $\partial f(\theta, j, k) / \partial \theta$.

We experienced considerable difficulty in obtaining positive definite covariance matrices from (B5) in some of our larger models because of the amount of noise in the Ω matrix, which has over 150,000 unique elements. Consequently, we set many elements of that matrix to zero and we smoothed many others with polynomial functions of age and year to obtain our standard error estimates.

Mappings. For Table 1 and related tables in the paper, f is just a linear regression function of age, year, and other variables corresponding to the location of the element in the covariance matrix. For the models presented in Tables 4 and 5, a mapping from the assumed error-components process to f is required. For the models in Table 4, which are independent of calendar time, the covariances (f) implies are easily obtained, but this is not the case for the models in Table 5 where the covariances are functions of calendar time. For those, we derive f at each age and year by recursively deriving successive variances and covariances over the life cycle for each cohort beginning at the beginning of the life cycle (age 20) and proceeding forward simultaneously in age and year. No initial conditions problems per se

are presented in our model because the life cycle has a finite start date and end date by assumption; no stationarity assumptions are made or required in either age or calendar time.

For brevity we will present the mapping only for the most complex model, that in equations (2)-(4). Letting $\delta_t = \text{Var}(\xi_{iat})$ and $\psi_t = \text{Var}(\omega_{iat})$, we have:

$$\text{Var}(y_{iat}) = \alpha_t^2 \text{Var}(\mu_{iat}) + \text{Var}(v_{iat}) \quad (\text{B6})$$

$$\begin{aligned} \text{Cov}(y_{iat}, y_{i,a-s,t-s}) &= \alpha_t \alpha_{t-s} \text{Cov}(\mu_{iat}, \mu_{i,a-s,t-s}) \\ &\quad + \text{Cov}(v_{iat}, v_{i,a-s,t-s}) \end{aligned} \quad (\text{B7})$$

$$\text{Var}(\mu_{i1t}) = \sigma_\mu^2 \quad (\text{B8})$$

$$\text{Var}(\mu_{iat}) = \text{Var}(\mu_{i,a-1,t-1}) + \psi_t, \quad a > 1 \quad (\text{B9})$$

$$\text{Cov}(\mu_{i2t}, \mu_{i1,t-1}) = \sigma_\mu^2 \quad (\text{B10})$$

$$\text{Cov}(\mu_{iat}, \mu_{i,a-s,t-s}) = \text{Cov}(\mu_{i,a-1,t-1}, \mu_{i,a-s,t-s}) \quad (\text{B11})$$

$$\text{Var}(v_{i1t}) = \delta_t \quad (\text{B12})$$

$$\text{Var}(v_{iat}) = \rho_t^2 \text{Var}(v_{i,a-1,t-1}) + \delta_t + \theta_t^2 \delta_t + 2\rho_t \theta_t \delta_{t-1}, \quad a > 1 \quad (\text{B13})$$

$$\text{Cov}(v_{iat}, v_{i,a-1,t-1}) = \rho_t \text{Var}(v_{i,a-1,t-1}) + \theta_t \delta_{t-1} \quad (\text{B14})$$

$$\begin{aligned} \text{Cov}(v_{iat}, v_{i,a-s,t-s}) &= \rho_t \text{Cov}(v_{i,a-1,t-1}, v_{i,a-s,t-s}) \\ &\quad + \theta_t \delta_{t-1}, \quad s > 1 \end{aligned} \quad (\text{B15})$$

Equations (B6) and (B7) provide recursion relationships for the variances and covariances, respectively, of equation (2) in the text. Equations (B8)-(B11) are recursion relationships for the variances and covariances of the permanent component, while (B12)-(B15) are recursion relationships for the variances and covariances of the transitory component. Conditional on values of the initial permanent variance (σ_μ^2) and the five parameters α_t , ρ_t , θ_t , ψ_t , and δ_t , all variances and covariances can be calculated by starting at $a=1$ for each cohort (each cohort begins at $a=1$ at a different calendar time, t) and by moving recursively forward over the life cycle using the formulae.

Appendix C

Relation of Mobility to Covariance Structure

Assume we have a random sample of n individuals with earnings observed at two points in time. We denote the earnings of individual i at time t as y_{it} ($i=1, \dots, n$; $t=1, 2$). Although earnings are independent across individuals we assume that they are correlated over time for the same individual. We also assume that the two earnings observations for each individual follow a bivariate normal distribution, with means zero and with $\text{Var}(y_{it}) = \sigma_y^2$ and $\text{Cov}(y_{it}, y_{it'}) = \rho\sigma_y^2$ ($t \neq t'$). Let P denote the probability that there are no changes in rank in the distribution from $t=1$ to $t=2$. We shall demonstrate that $\partial P / \partial \rho > 0$ and that P is independent of σ_y^2 .

Ordering the individuals from $i=1$ to $i=n$ by rank, we have:

$$P = n! \text{ Prob}(y_{11} < y_{21} < y_{31} < \dots < y_{n1}, y_{12} < y_{22} < y_{32} < \dots < y_{n2}) \quad (C1)$$

since there are $n!$ possible orderings of the n individuals, each ordering with the same probability.

It is sufficient to compare only the change in relative rank for any given pair of individuals i and j , since the result will generalize to all pairs. Let

$$\begin{aligned} P' &= \text{Prob}(y_{i1} < y_{j1}, y_{i2} < y_{j2}) + \text{Prob}(y_{i1} > y_{j1}, y_{i2} > y_{j2}) \\ &= 2 * \text{Prob}(y_{i1} < y_{j1}, y_{i2} < y_{j2}) \end{aligned} \quad (C2)$$

Defining

$$w_1 = y_{i1} - y_{j1} \quad (C3)$$

$$w_2 = y_{i2} - y_{j2} \quad (C4)$$

we have that w_1 and w_2 are distributed bivariate normal with means zero and

with $\text{Var}(w_t) = 2\sigma_Y^2$ and $\text{Cov}(w_1, w_2) = 2\sigma_Y^2\rho$. Hence

$$P' = \int_{-\infty}^0 \int_{-\infty}^0 (2\sigma_Y^2)^{-1} b(w_1, w_2; \rho) dw_1 dw_2 \quad (C5)$$

$$= \int_{-\infty}^0 \int_{-\infty}^0 b(\hat{w}_1, \hat{w}_2; \rho) d\hat{w}_1 d\hat{w}_2 \quad (C6)$$

where b is the unit bivariate normal density and where $\hat{w}_j = w_j/\sqrt{2}\sigma_Y$. Thus P' is only a function of ρ and not a function of σ_Y^2 .

The partial derivative of a bivariate normal cumulative distribution function w.r.t. ρ is equal to the bivariate density evaluated at the upper limits. Hence $\partial P'/\partial \rho = b(0, 0; \rho) > 0$.

Notes

1. In prior work, we have treated each wave of the PSID as an independent cross section and we have compared trends in earnings differentials to those in the CPS (Gottschalk and Moffitt, 1992). We found overall conformity of the direction of the trends in the two data sets in both within-group and between-group earnings differentials.
2. Attrition in the PSID has reached approximately 50 percent by 1988, and therefore a continuous-sample restriction would severely reduce the sample size and hinder the analysis. Fortunately, despite this heavy attrition, there is little evidence of significant attrition bias in the PSID in the studies that have examined it to date (Beckett et al., 1988; Fitzgerald, Gottschalk, and Moffitt, forthcoming).
3. We delete the top and bottom one-percent of the earnings and wage observations within each age-year covariance cell. These outliers introduce noise into the trends in variances and covariances. This trimming also eliminates the top-coded earnings observations in the PSID. Results on untrimmed data show the same patterns as those we present below but with larger standard errors.
4. The regression coefficients are available upon request.
5. In addition, if the transitory component is serially correlated, the variance of average earnings will be affected by the variance of the transitory component as well. The method we use in this paper confirms the crude estimate in Gottschalk and Moffitt (1994).
6. The variances and covariances are averaged over the age groups in question and over the single-order lags (e.g., over the first, second, third, and fourth-order lags for the plot "lag 1-4").
7. The difference in the growth rates of high-order and low-order covariances, and their different implications for the growth of the transitory variance, will be discussed further below.
8. However, variances are more procyclical than covariances (as can also be seen in Figures 1(a)-(d)), an indirect indication that transitory variances are more procyclical than "permanent" variances. We should note that the unemployment rates are detrended.
9. That the two models fit approximately the same is not surprising since the random growth model implies variances rising over the life cycle with the square of age whereas the random walk model implies linearly increasing variances. There is relatively little curvature in the life cycle profile of variances in our data, which is no doubt the reason for the similarity of fit. Abowd and Card (1989) examined fourth moments of the data and found the random walk model to provide a superior fit than the random growth model because earnings differences were weakly correlated over time.
10. We make no attempt to explicitly identify measurement error components although such error will unquestionably enter in various places. Although classical measurement error could be captured by the innovations in the MA process, more recent work on error in earnings reports suggests that measurement error is serially correlated (Bound et al., 1990; Bound and Krueger, 1991). Hence the parameters θ_t and ρ_t could pick up some measurement

error as well. Unfortunately, the Bound-Krueger and Bound et al. studies only had two periods of validated earnings data, and hence AR and MA components of measurement error could not be identified.

11. At the third lag, for example, the covariance of differences is reduced by ρ raised to the sixth power, which in our case ($\rho=.662$) is .084. It falls further at higher lags. Although it is clear from Table 1 that our data require a first-order autoregressive parameter to explain the decline in autocovariances after the second lag, the magnitude may be too small for statistical significance when estimation is conducted in differences. Another possible explanation for this minor difference is that we have a panel of 19 periods, whereas MaCurdy and Abowd-Card only had panels of 10 and 11, respectively; our longer panel may give us more power in detecting small autoregressive influences at longer lags.

12. The chi-squared statistics for both specifications are far above the one-percent critical values of 600-624, implying considerable unexplained variation in the model. However, replacing the time trends with year dummies, whose results will be discussed momentarily, reduce the chi-squareds to 739.

13. The order of parameter change in Table 6 does not materially affect these conclusions because the permanent and transitory components are additive in the total variance and hence do not interact. For example, introducing the change in the variance of ξ first increases the four variances from their 1969 values to .277, .305, .317, and .328 for the four respective ages shown in the Table.

14. The change induced by the trend coefficient for the variance of the random walk is negligible in magnitude and hence is not broken out separately; the changes in the last row are entirely due to the change in ρ .

15. An alternative computation is to compute the "steady-state" variances implied by the values of the parameters in 1987, and to compare these to the 1969 steady-state variances shown in the first row of Table 6. The 1987 variances in Table 6 reflect the historical experience of the shocks from 1969 to 1987 and, because of the autoregressive structure of the model, do not represent the steady-state values. However, because the autoregressive lag in the variances is so short--of negligible importance after three years--the 1987 steady-state variances differ from those in the last row of Table 6 only at the second decimal place.

16. The increase in transitory variance could in principle be the result of increasing measurement error in the PSID, but there is no evidence that it did so or any reason to think it would have increased more in the 1980s. Nor is there any evidence that there has been a change in the accuracy of earnings data in the PSID. The fraction of earnings observations that are imputed, combining what the PSID calls "major" and "minor" imputations, is only 1.6 percent in our sample averaged over all years. This low percent probably reflects better reporting among prime-age white males than other population groups. More important, the fraction has not changed over the period, varying only between .6 percent and 2.1 percent, with a slight downward trend over time. As a consequence, estimates of the model shown in Table 6 change only at the third or fourth decimal place when imputed earnings observations are deleted. In addition, there has been no change in the coding procedures used to detect "erroneous" earnings. Those procedures are documented for coders, and the same documents have been used for the entire PSID.

17. Over the entire lifetime, for example, an increase in the permanent variance must both increase the variance of lifetime earnings and lower mobility between average earnings in the first part of the life cycle and in

the second part, at least as long as transitory components die out within those parts.

18. The limited number of observations in the sample prevent us from disaggregating the quantiles further. Some prior analyses in the literature have been able to use a finer set of quantiles by pooling the data across years and across ages. However, not only have we found pooling across years to be incorrect, both our descriptive and error components analysis showed the necessity of conditioning on age.

19. This is to be expected since persons in the upper (lower) quintiles can only move down (up), whereas persons in the middle quintiles can move in either direction.

20. Given our age grouping, "35" stands for 30-39, "36" stands for 31-40, and "40" stands for 35-44.

21. These trends by themselves have no necessary implication for our prior results. Mean weeks of work have fallen in our sample whether these zero-weeks observations are included or not (although they have fallen more when the zeros are included than when they are not). More important, a change in mean weeks worked, negative or positive, has no necessary implication for changes in variances.

22. For the descriptive regressions, the off-diagonal and diagonal coefficients are 1.296 and 1.109, respectively, both significant at the 10 percent level. The magnitudes are considerably larger than those for conditional weeks worked because variances and covariances showed larger absolute increases over the period. Estimates of the error components model reduce the magnitude and significance level of the trend in the transitory component, and increase them for the permanent component.

23. A detailed comparison of the PSID and CPS in this dimension can be found in our prior benchmarking exercise (Gottschalk and Moffitt, 1992).

24. The sum of the permanent and transitory components of the within values are almost identical to the values given in Table 6. As noted previously, the steady-state predicted values from the model are essentially identical to the values predicted historically because the latter only include the influence of "history"--that is, the fact that variances have been growing over time and hence have not been at their steady-state value for the whole period--but history is unimportant after approximately three years.

25. To some extent the small magnitude of the change induced by including the between simply reflects the relatively small R-squareds in all log earnings regressions when only education and age are the explanators; hence trends in the covariance structure of total log earnings are dominated by trends in the covariance structure of the residuals. Note as well that the correlation coefficients in Table 11 imply even less reduction in mobility than was found for the within analysis. The trends in the correlation coefficients arise from the size of the proportionate increase in the permanent variance, not its absolute size, and the proportionate increase in that variance is smaller for the between than for the within.

26. The individuals in age group 50-59 cannot be followed at all. However, their earnings are used to construct diagonal elements of the covariance matrix (i.e., variances rather than covariances).

27. To compute the covariances between those who are, say, 20-29 in 1970 and (hence) 30-39 in 1980, we use only those who are present in both years.

28. Abowd and Card (1989) also used the identity matrix because of problems similar to those discussed by Altonji and Segal.

29. Each individual in our data set contributes to only a subset of the moments in Ω (see Appendix A); we do not adjust the notation in (B4) for this, but instead leave it implicit.

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Log Earnings Variances and Covariances by Year and by Lag Order

Figure 1(a): All Ages

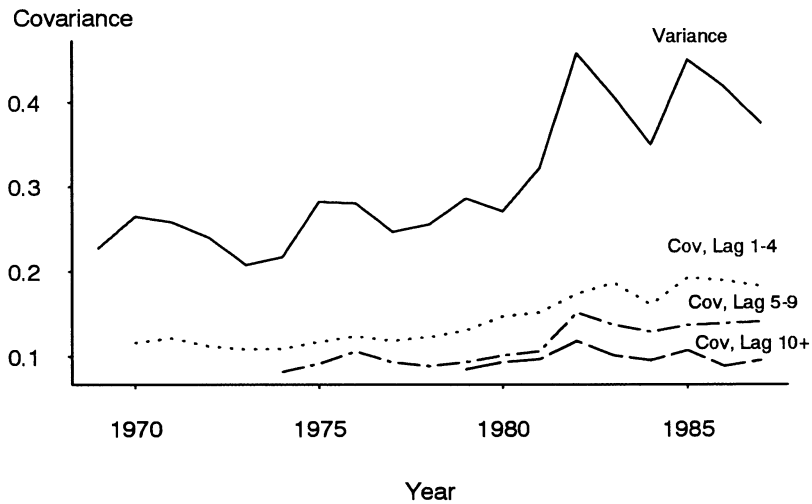


Figure 1(b): Ages 25-34

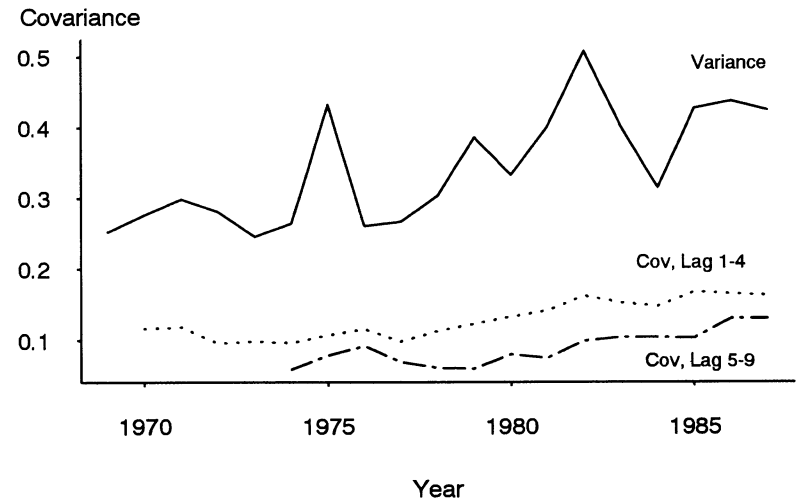


Figure 1(c): Ages 35-44

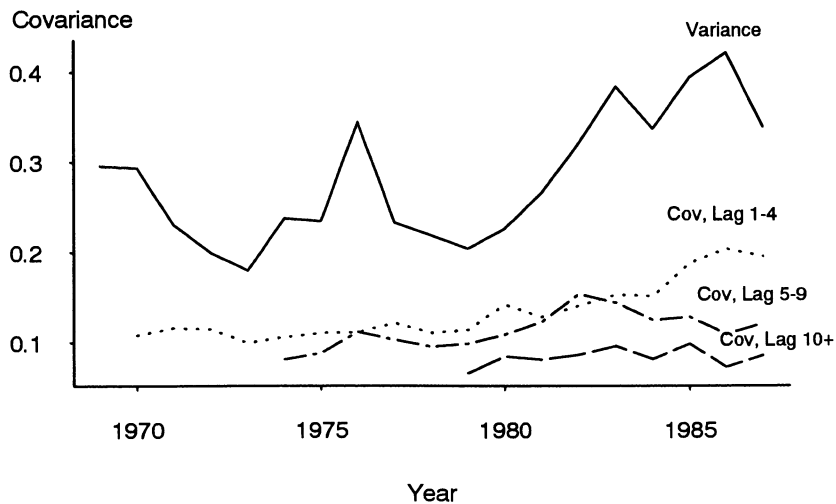


Figure 1(d): Ages 45-54

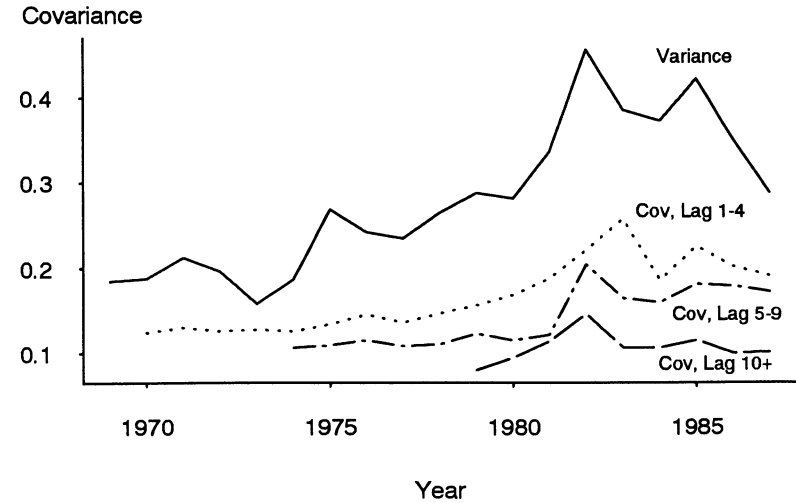


Figure 2: Estimated Trends in Permanent and Transitory Components

Figure 2(a)

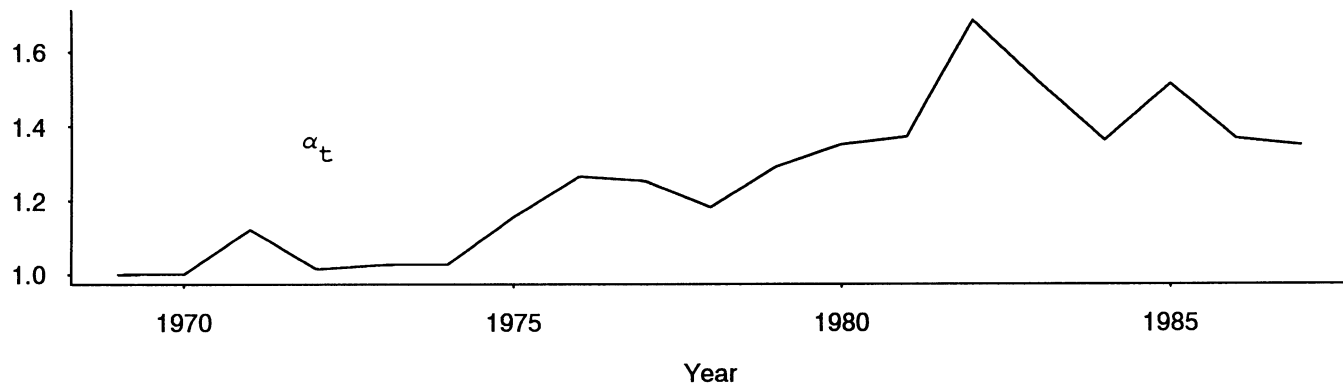


Figure 2(b)

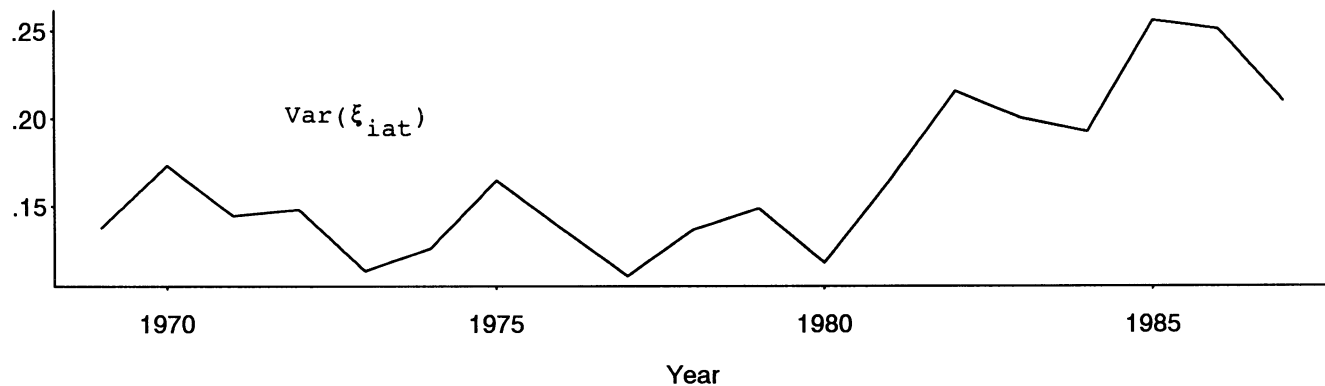


Figure 3: Mobility Rates and Correlation Coefficients by Year and Age

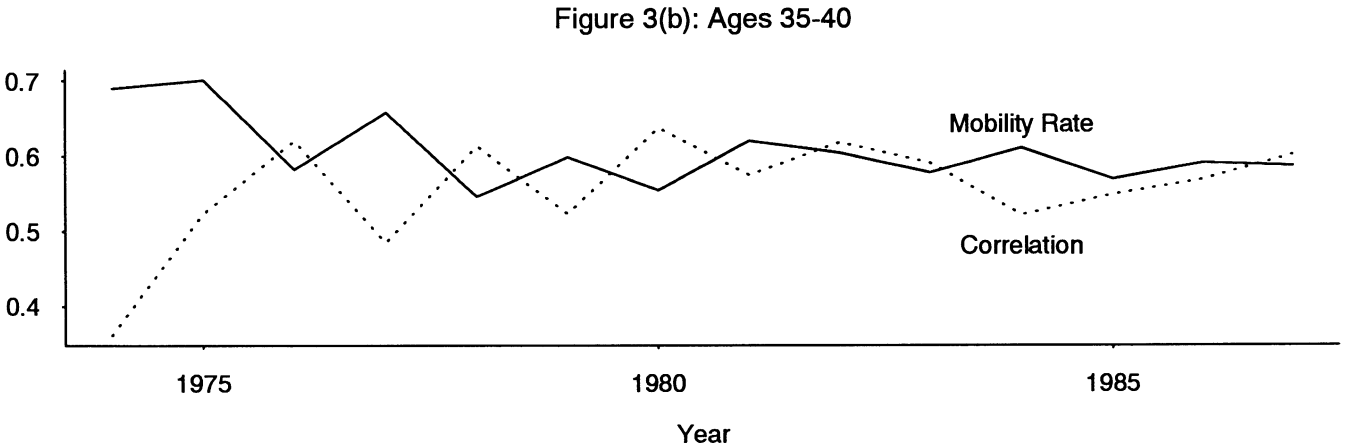
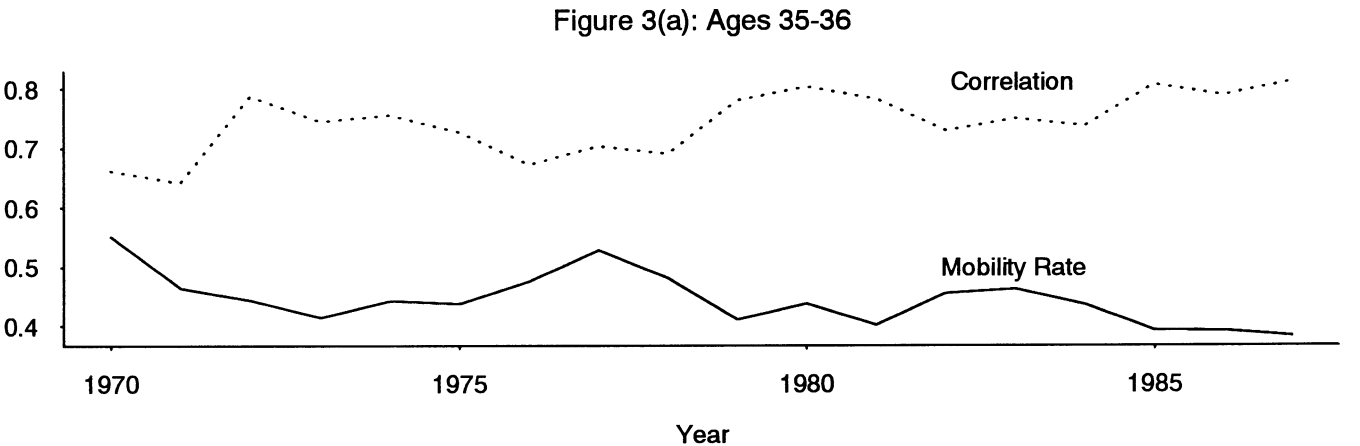


Table 1
Descriptive Covariance Regressions for Log Annual Earnings

	(1)	(2)	(3)	(4)
Intercept	.1253* (.0006)	.0846* (.0045)	.0823* (.0049)	.0878* (.0082)
D	.1811* (.0050)	.1296* (.0080)	.1319* (.0087)	.1656* (.0132)
t	-	.0058* (.0003)	.0061* (.0003)	.0059* (.0003)
Dt	-	.0059* (.0005)	.0056* (.0005)	.0014 (.0012)
A ₂	-	.0023* (.0003)	.0023* (.0003)	.0012 (.0004)
DA ₂	-	-.0029* (.0003)	-.0029* (.0003)	-.0038* (.0005)
(A ₂ -A ₁)	-	-.0165* (.0014)	-.0168* (.0014)	-.0135* (.0015)
(A ₂ -A ₁) ² /100	-	.0476* (.0068)	.0537* (.0068)	.0712* (.0085)
U ₂	-	-	.0065* (.0010)	.0059* (.0009)
DU ₂	-	-	.0097* (.0027)	.0101* (.0027)
U ₁	-	-	.0020 (.0011)	.0021* (.0011)
tA ₂ /10	-	-	-	.0010* (.0003)
tDA ₂ /10	-	-	-	.0012* (.0005)
t(A ₂ -A ₁)/10	-	-	-	-.0042* (.0011)
Chi-squared ^a (df)	2272 (551)	1029 (545)	1138 (542)	1101 (539)

Notes: Standard errors in parentheses
*: significant at the 10 percent level
n=553
Unemployment rate is for all U.S. male civilians 20 and over.
D=diagonal dummy; A₂ = the older age minus 20; A₁ = the younger age minus 20; t = year at age A₂ minus 1969; U₂ = unemployment rate at age A₂; U₁ = unemployment rate at age A₁
^a Statistic= $ne'S^{-1}e$, where e is the vector of estimated residuals and S is their empirical covariance matrix.

Table 2
Log Annual Earnings Covariance Regressions
by Lag Order and Time Period:
Selected Coefficients

	All Years	1969-1980	1981-1987
<u>All Lag Orders</u>			
t	.0061* (.0004)	.0056* (.0008)	.0066* (.0008)
Dt	.0056* (.0008)	-.0008 (.0014)	.0183* (.0022)
<u>Lag Orders 1-3</u>			
t	.0066* (.0007)	.0049* (.0013)	.0093* (.0017)
Dt	.0051* (.0011)	-.0001 (.0019)	.0156* (.0031)

Notes: Standard errors in parentheses
*: significant at the 10 percent level
Acronyms: see Table 1
Also included in regressions: $D, A_2, DA_2, A_2 - A_1, (A_2 - A_1)^2, U_2, DU_2, U_1$

Table 3
Log Annual Earnings Covariances and Correlations by Age
(pooled over years)

Leads	Covariance				Correlation			
	20- 29	30- 39	40- 49	50- 59	20- 29	30- 39	40- 49	50- 59
0	.35 (.009)	.30 (.008)	.27 (.008)	.32 (.011)	1.0	1.0	1.0	1.0
1	.17 (.005)	.18 (.005)	.18 (.005)	-	.63	.74	.80	-
2	.13 (.005)	.15 (.004)	.17 (.005)	-	.51	.64	.73	-
3	.12 (.005)	.14 (.004)	.15 (.005)	-	.43	.59	.69	-
4	.11 (.005)	.13 (.005)	.15 (.005)	-	.40	.55	.66	-
5	.11 (.006)	.12 (.005)	.15 (.006)	-	.37	.50	.65	-
6	.09 (.005)	.11 (.005)	.15 (.007)	-	.34	.49	.64	-
7	.09 (.005)	.12 (.006)	.14 (.007)	-	.32	.51	.61	-
8	.10 (.006)	.12 (.006)	.13 (.006)	-	.35	.49	.58	-
9	.09 (.006)	.11 (.006)	.12 (.006)	-	.32	.46	.53	-
10	.09 (.007)	.11 (.007)	.12 (.006)	-	.34	.47	.49	-
11	.07 (.007)	.12 (.009)	-	-	.29	.46	-	-
12	.09 (.010)	.11 (.008)	-	-	.29	.47	-	-
13	.09 (.012)	.10 (.008)	-	-	.29	.48	-	-
14	.08 (.011)	.10 (.007)	-	-	.28	.43	-	-
15	.10 (.013)	.10 (.009)	-	-	.32	.45	-	-

Table 3

(cont'd)

Leads	Covariance				Correlation			
	20- 29	30- 39	40- 49	50- 59	20- 29	30- 39	40- 49	50- 59
16	.09 (.015)	.10 (.010)	-	-	.29	.42	-	-
17	.09 (.018)	.11 (.013)	-	-	.30	.45	-	-
18	.10 (.025)	.11 (.019)	-	-	.37	.37	-	-

Notes: Cell sample sizes range from 183 to 7358.
Standard errors in parentheses

Table 4
Error Components Models for Log Real Annual Earnings
(no calendar time effects)

	RE + AR(1)	RE + ARMA(1,1)	Random-Walk RE + ARMA(1,1)	Random-Growth RE + ARMA(1,1)
	(I)	(II)	(III)	(IV)
σ_{μ}^2	.114 (.008)	.070 (.012)	.090 (.015)	.092 (.005)
ρ	.354 (.017)	.906 (.026)	.622 (.025)	.657 (.069)
σ_{η}^2	.167 (.013)	-	-	-
θ	-	-.670 (.017)	-.344 (.045)	-.362 (.092)
σ_{ξ}^2	-	.188 (.013)	.169 (.016)	.166 (.007)
σ_{ω}^2	-	-	.159 ^a (.025)	-
σ_{ϕ}^2	-	-	-	.462 ^a (.069)
Chi-squared (df)	2388 (550)	2262 (549)	2094 (548)	2080 (548)

Notes: Standard errors in parentheses
All coefficients significant at 10 percent level
RE=random effect
^a Parameter multiplied by 100

Model I: (1) $e_{ia} = \mu_i + v_{ia}$
(2) $v_{ia} = \rho v_{i,a-1} + \eta_{ia}$

Model II: (1), (2), plus
(3) $\eta_{ia} = \xi_{ia} + \theta \xi_{i,a-1}$

Model III: (2), (3), plus
(1') $e_{ia} = \mu_{ia} + v_{ia}$
(4) $\mu_{ia} = \mu_{i,a-1} + \omega_{ia}$
with $\sigma_{\mu}^2 = \text{Var}(\mu_{i1})$

Model IV: (2), (3), plus
(1'') $e_{ia} = \mu_i + a\phi_i + v_{ia}$

All primary error components assumed independent.

Table 5
Error Components Models for Log Annual Earnings
with Calendar Time Effects

	(1)	(2)
<hr/>		
α_t :		
Year	.029* (.003)	.023* (.006)
$\text{Var}(\xi_{iat})$:		
Year	.005* (.001)	.005* (.001)
Constant	.117* (.009)	.118* (.010)
ρ_t :		
Year	-	.008 (.015)
Constant	.641* (.063)	.578* (.191)
θ_t :		
Year	-	-.003 (.018)
Constant	-.367* (.078)	-.352* (.227)
$\text{Var}(\omega_{iat})$: ^a		
Year	-	.001 (.010)
Constant	.100* (.013)	.110* (.017)
σ_μ^2	.056* (.003)	.061* (.006)
Chi-Squared (df)	1082 (546)	1076 (543)

Notes: Standard errors in parentheses
*: significant at 10 percent level
Year=0 in 1969, =1 in 1970, etc.
^aParameters multiplied by 100.

Table 6

Effects of Parameter Changes on Log Annual Earnings Variances,
1969-1987

	Variances by Age			
	20	30	40	50
1969 Values of all Parameters	.179	.199	.210	.221
1987 Value of α_t only	.240	.273	.296	.318
1987 Values of α_t and Variance of ξ_{iat}	.338	.379	.401	.423
1987 Values of α_t , Variance of ξ_{iat} , and θ_t	.338	.372	.394	.416
1987 Values of all Parameters	.338	.401	.423	.445

Table 7
One-Year Quintile Mobility Rates for Log Annual Earnings:
All Years and Ages

Quintile at t-1	Sum	Quintile Distribution at t				
		Bottom Fifth	Next to Bottom Fifth	Middle Fifth	Next to Top Fifth	Top Fifth
Bottom Fifth	100	67	21	8	3	1
Next to Bottom Fifth	100	20	49	22	7	2
Middle Fifth	100	7	21	44	22	6
Next to Top Fifth	100	4	7	20	47	22
Top Fifth	100	2	3	7	20	69

Notes: 477 observations per row

Table 8
Year Coefficients in Quintile Mobility Regressions

	All	Initial Quintile Location				
		Bottom Fifth	Next to Bottom Fifth	Middle Fifth	Next to Top Fifth	Top Fifth
<u>All Lag Orders</u>						
Year	-.0020* (.0004)	-.0040* (.0008)	-.0028* (.0008)	-.0010 (.0007)	-.0007 (.0007)	-.0020* (.0007)
<u>Lag Orders 1-4</u>						
1969-1980	-.0013 (.0011)	-.0051* (.0018)	-.0008 (.0019)	.0001 (.0019)	.0006 (.0017)	-.0017 (.0017)
1981-1987	-.0025* (.0014)	-.0022 (.0023)	-.0038 (.0024)	-.0026 (.0024)	-.0017 (.0021)	-.0026 (.0022)
<u>Lag Orders 5+</u>						
1969-1980	-.0067* (.0017)	-.0083* (.0035)	-.0035 (.0034)	-.0014 (.0030)	-.0087* (.0034)	-.0108* (.0034)
1981-1987	-.0005 (.0009)	-.0027 (.0018)	-.0044* (.0018)	-.0004 (.0015)	.0021 (.0018)	.0020 (.0018)

Notes: Standard errors in parentheses
*: significant at the 10 percent level
Sample sizes are 477 for all-lag-order sample, and 198 and 279 for the 1-4 lag-order and 5+ lag-order samples, respectively.
Dependent variable: fraction of population in the relevant age-year-quintile cell that changed quintiles over the lag orders shown.
Independent variables in addition to time trends: A_2 , $A_2 - A_1$, and $(A_2 - A_1)$ squared (see notes to Table 2 for definitions).¹

Table 9

Covariance Analysis of Real Weekly Wages
and Annual Weeks of Work

	Log Real Weekly Wage	Log Annual Weeks Worked	Annual Weeks Worked ^a
<u>Descriptive Regressions</u> ^b			
t	.0044* (.0004)	.0008* (.0002)	.485* (.110)
Dt	.0043* (.0008)	.0021* (.0004)	.846* (.215)
<u>Error Components Model</u> ^c			
α_t	.024* (.002)	.075 (.065)	.013 (.014)
Var(ξ_{iat})	.003* (.001)	.002* (.000)	.487* (.283)

Notes: Standard errors in parentheses

*: significant at the 10 percent level

^aFor positive weeks of work^bFor specification in third column of Table 2^cTrend coefficients. For specification in first column of Table 5

Table 10
Steady-State Variance Components
Implied by the Estimated Error Components Models

	Within			Total		
	Permanent	Transitory	Rho	Permanent	Transitory	Rho
<u>Age 20</u>						
1969	.061	.118	.34	.088	.118	.43
1987	.122	.216	.36	.154	.212	.42
<u>Age 30</u>						
1969	.072	.127	.36	.109	.127	.46
1987	.144	.258	.36	.192	.256	.43

Notes: The permanent variance is the estimated value of $\text{Var}(\mu_{iat})$ at 1969 and 1987 values. The transitory variance is the estimated value of $\text{Var}(v_{iat})$ at 1969 and 1987 values of the parameters. The model estimates in column (1) of Table 5 and the analogous estimates for total log earnings are used.