

### Problem Set 1: Solutions

1. (a)  $f(x)$  is not nonnegative over the stated range. For example,  $f(2) = -4$ . So it cannot be a pdf.

(b)  $f(x)$  is nonnegative over the stated range. If it integrates to 1, then  $f(x)$  is a pdf.

$$f(x) = C(2x - x^2) \Rightarrow \int_0^{1.5} C(2x - x^2) = 1 \Rightarrow C \int_0^{1.5} (2x - x^2) = 1 \Rightarrow C \left[ x^2 - \frac{x^3}{3} \right]_0^{1.5} = C \left( 1.5^2 - \frac{1.5^3}{3} \right) = 1.125C$$

$$\therefore C = 8/9$$

2. Event A: The first statement is true

Event B: The second islander says that it is true.

By Bayes Rule

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

$$P(B) = P(B|A)P(A) + P(B|\bar{A})P(\bar{A})$$

$$\therefore P(A|B) = \frac{(1/3) * (1/3)}{(1/3) * (1/3) + (2/3) * (2/3)} = \frac{1}{5}$$

$$3.(i) F_X(x) = \frac{x+1}{2}$$

$\log(X+1)$  is monotone increasing.

$$g^{-1}(\cdot) = \exp(\cdot) - 1$$

$$\therefore F_Y(y) = \frac{\exp(y) - 1 + 1}{2} = e^y / 2$$

$$f_Y(y) = e^y / 2$$

Support is from  $-\infty$  to  $\log(2)$ .

$$\int_{-\infty}^{\log(2)} \frac{e^y}{2} dy = \left[ \frac{e^y}{2} \right]_{-\infty}^{\log(2)} = \frac{e^{\log(2)}}{2} = \frac{2}{2} = 1$$

(ii)  $|X|$  is not monotone.

$$F_Y(y) = P(|X| \leq y) = F_X(y) - F_X(-y) = \frac{y+1}{2} - \frac{-y+1}{2} = y$$

$$\therefore f_Y(y) = 1$$

Support is from 0 to 1.

$$\int_0^1 1 dy = [y]_0^1 = 1$$

(iii)  $X^3$  is monotone increasing.

$$g^{-1}(\cdot) = (\cdot)^{1/3}$$

$$\therefore F_Y(y) = \frac{y^{1/3} + 1}{2}$$

$$\therefore f_Y(y) = \frac{1}{2} \frac{1}{3} y^{-2/3} = \frac{1}{6y^{2/3}}$$

Support is from -1 to +1.

$$\int_{-1}^1 \frac{1}{6y^{2/3}} dy = \left[ \frac{1}{2} y^{1/3} \right]_{-1}^1 = \frac{1}{2} + \frac{1}{2} = 1$$

$$4. p(i) = 0.5(p(i+1) + p(i-1))$$

$$p(0) = 0 \text{ and } p(15) = 1$$

I want to prove  $(i+1)p(i) = ip(i+1)$  by induction.

For  $i = 0$ ,  $p(0) = 0$  and so this works.

Suppose that  $(i+1)p(i) = ip(i+1)$

$$p(i+1) = 0.5(p(i+2) + p(i))$$

$$\therefore p(i+1) = 0.5(p(i+2) + \frac{i}{i+1} p(i+1))$$

$$\therefore 2p(i+1) = p(i+2) + \frac{i}{i+1} p(i+1)$$

$$\therefore 2p(i+1)(i+1) = (i+1)p(i+2) + ip(i+1)$$

$$\therefore [2i+2-i]p(i+1) = (i+1)p(i+2)$$

$$\therefore (i+2)p(i+1) = (i+1)p(i+2)$$

and so the relation has been shown by proof by induction.

$$p(15) = 1. \text{ Working backwards, } p(n) = \prod_{j=n}^{14} \frac{j}{j+1} = \frac{n}{15}.$$

So  $p(5) = 1/3$  is the probability that A wins starting with \$5.

$$5. E(X^2) = \int_{-\infty}^{\infty} x^2 f(x) dx = \int_{-\infty}^{\infty} x^2 \frac{1}{\sqrt{2\sigma}} \exp\left(-\frac{|x-\mu|}{\sigma/\sqrt{2}}\right) dx$$

$$= \int_{-\infty}^{\mu} x^2 \frac{1}{\sqrt{2\sigma}} \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) dx + \int_{\mu}^{\infty} x^2 \frac{1}{\sqrt{2\sigma}} \exp\left(\frac{\mu-x}{\sigma/\sqrt{2}}\right) dx$$

Integration by parts

$$\int_a^b f(x)g'(x)dx = [f(x)g(x)]_a^b - \int_a^b f'(x)g(x)dx$$

$$\text{Consider } \int_{-\infty}^{\mu} x^2 \frac{1}{\sqrt{2\sigma}} \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) dx.$$

$$\text{Let } f(x) = \frac{x^2}{\sqrt{2\sigma}} \text{ and } g(x) = \frac{\sigma}{\sqrt{2}} \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) \Rightarrow g'(x) = \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right).$$

$$\therefore \int_{-\infty}^{\mu} \frac{x^2}{\sqrt{2\sigma}} \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) dx = \left[ \frac{x^2}{\sqrt{2\sigma}} \frac{\sigma}{\sqrt{2}} \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) \right]_{-\infty}^{\mu} - \int_{-\infty}^{\mu} \frac{2x}{\sqrt{2\sigma}} \frac{\sigma}{\sqrt{2}} \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) dx$$

$$= \left[ \frac{x^2}{2} \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) \right]_{-\infty}^{\mu} - \int_{-\infty}^{\mu} x \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) dx = \frac{\mu^2}{2} - \int_{-\infty}^{\mu} x \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) dx$$

By another application of integration by parts,

$$\begin{aligned} \int_{-\infty}^{\mu} x \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) dx &= \left[ x \frac{\sigma}{\sqrt{2}} \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) \right]_{-\infty}^{\mu} - \int_{-\infty}^{\mu} \frac{\sigma}{\sqrt{2}} \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) dx \\ &= \frac{\mu\sigma}{\sqrt{2}} - \left[ \frac{\sigma^2}{2} \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) \right]_{-\infty}^{\mu} = \frac{\mu\sigma}{\sqrt{2}} - \frac{\sigma^2}{2} \end{aligned}$$

$$\therefore \int_{-\infty}^{\mu} \frac{x^2}{\sqrt{2}\sigma} \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) dx = \frac{\mu^2}{2} - \frac{\mu\sigma}{\sqrt{2}} + \frac{\sigma^2}{2}$$

Now consider  $\int_{\mu}^{\infty} x^2 \frac{1}{\sqrt{2}\sigma} \exp\left(\frac{\mu-x}{\sigma/\sqrt{2}}\right) dx$ . Proceeding similarly,

$$\begin{aligned} \therefore \int_{\mu}^{\infty} \frac{x^2}{\sqrt{2}\sigma} \exp\left(\frac{\mu-x}{\sigma/\sqrt{2}}\right) dx &= \left[ -\frac{x^2}{\sqrt{2}\sigma} \frac{\sigma}{\sqrt{2}} \exp\left(\frac{\mu-x}{\sigma/\sqrt{2}}\right) \right]_{\mu}^{\infty} - \int_{\mu}^{\infty} \left\{ -\frac{2x}{\sqrt{2}\sigma} \frac{\sigma}{\sqrt{2}} \exp\left(\frac{\mu-x}{\sigma/\sqrt{2}}\right) \right\} dx \\ &= \left[ -\frac{x^2}{2} \exp\left(\frac{\mu-x}{\sigma/\sqrt{2}}\right) \right]_{\mu}^{\infty} - \int_{\mu}^{\infty} \{-x \exp\left(\frac{\mu-x}{\sigma/\sqrt{2}}\right)\} dx = \frac{\mu^2}{2} - \int_{\mu}^{\infty} \{-x \exp\left(\frac{\mu-x}{\sigma/\sqrt{2}}\right)\} dx \end{aligned}$$

Again, using integration by parts,

$$\begin{aligned} \int_{\mu}^{\infty} \{-x \exp\left(\frac{\mu-x}{\sigma/\sqrt{2}}\right)\} dx &= \left\{ \left[ x \frac{\sigma}{\sqrt{2}} \exp\left(\frac{\mu-x}{\sigma/\sqrt{2}}\right) \right]_{\mu}^{\infty} - \int_{\mu}^{\infty} \frac{\sigma}{\sqrt{2}} \exp\left(\frac{\mu-x}{\sigma/\sqrt{2}}\right) dx \right\} \\ &= -\frac{\mu\sigma}{\sqrt{2}} - \left[ -\frac{\sigma^2}{2} \exp\left(\frac{\mu-x}{\sigma/\sqrt{2}}\right) \right]_{\mu}^{\infty} = -\frac{\mu\sigma}{\sqrt{2}} - \frac{\sigma^2}{2} \end{aligned}$$

$$\therefore \int_{\mu}^{\infty} \frac{x^2}{\sqrt{2}\sigma} \exp\left(\frac{x-\mu}{\sigma/\sqrt{2}}\right) dx = \frac{\mu^2}{2} + \frac{\mu\sigma}{\sqrt{2}} + \frac{\sigma^2}{2}$$

Adding the pieces together,

$$\therefore E(X^2) = \mu^2 + \sigma^2$$

$$\therefore \text{Var}(X) = E(X^2) - E(X)^2 = \mu^2 + \sigma^2 - \mu^2 = \sigma^2$$

$$6. P(X = x) = \sum_{x=0}^n \frac{n!}{x!(n-x)!} p^x (1-p)^{n-x}$$

$$m(t) = E(e^{tx}) = \sum_{x=0}^n e^{tx} \frac{n!}{x!(n-x)!} p^x (1-p)^{n-x} = \sum_{x=0}^n \frac{n!}{x!(n-x)!} [pe^t]^x (1-p)^{n-x}$$

$$\therefore m(t) = (pe^t + 1 - p)^n$$

$$m'(t) = n(pe^t + 1 - p)^{n-1} pe^t$$

$$\therefore m'(0) = np$$

$$m''(t) = (pe^t + 1 - p)^{n-1} npe^t + (n-1)(pe^t + 1 - p)^{n-2} pe^t npe^t$$

$$\therefore m''(0) = np + (n-1)np^2$$

$$Var(X) = m''(0) - m'(0)^2 = np + (n-1)np^2 - n^2 p^2 = np + n^2 p^2 - np^2 - n^2 p^2 = np - np^2 = np(1-p)$$

7. Matlab code.

```
rand('seed',123);
x=randn(10000,1);
y=-log(x);
hist(y);
z=[0.1:0.1:12]';
hold on;
plot(z,10000*exp(-z),'Linewidth',2);
axis([0 12 0 10000]);
```

The theoretical pdf is  $f(y) = e^{-y}$ .

Resulting graph:

